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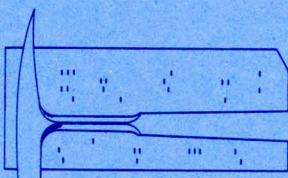
**FORTRAN IV  
PROGRAM FOR THE  
GENERALIZED STATISTICAL  
DISTANCE AND ANALYSIS  
OF COVARIANCE MATRICES  
FOR THE  
CDC 3600 COMPUTER**

**By**

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in cooperation with the  
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The "FORTRAN IV program for the generalized statistical distance and analysis of covariance matrices for the CDC 3600 computer", by R.A. Reyment, Hans-Ake Ramden, and W.J. Wahlstedt will be made available on magnetic tape for a limited time for \$25.00 (U.S.). An extra charge of \$10.00 will be made if punched cards are required.

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## COMPUTER CONTRIBUTIONS

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### Computer Contribution

- |   |                |
|---|----------------|
| 1. Mathematical simulation of marine sedimentation with IBM 7090/7094 computers, by J.W. Harbaugh, 1966   | (out of print) |
| 2. A generalized two-dimensional regression procedure, by J.R. Dempsey, 1966  | \$0.50         |
| 3. FORTRAN IV and MAP program for computation and plotting of trend surfaces for degrees 1 through 6, by Mont O'Leary, R.H. Lippert, and O.T. Spitz, 1966 | \$0.75         |
| 4. FORTRAN II program for multivariate discriminant analysis using an IBM 1620 computer, by J.C. Davis and R.J. Sampson, 1966                             | \$0.50         |
| 5. FORTRAN IV program using double Fourier series for surface fitting of irregularly spaced data, by W.R. James, 1966                                     | \$0.75         |
| 6. FORTRAN IV program for estimation of cladistic relationships using the IBM 7040, by R.L. Bartcher, 1966  | \$1.00         |
| 7. Computer applications in the earth sciences: Colloquium on classification procedures, edited by D.F. Merriam, 1966                                     | \$1.00         |
| 8. Prediction of the performance of a solution gas drive reservoir by Muskat's Equation, by Apolonio Baca, 1967   | \$1.00         |
- (continued on inside back cover)

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# FORTRAN IV PROGRAM FOR GENERALIZED STATISTICAL DISTANCE AND ANALYSIS OF CO-VARIANCE MATRICES FOR THE CDC 3600 COMPUTER

by

R.A. Reyment, Hans-Ake Ramden and Warren J. Wahlstedt

## ABSTRACT

Heterogeneity in sample covariance matrices, deriving from differences in the orientation of major axes of the ellipsoids of scatter, may be of common occurrence. The generalized test of equality of covariance matrices will give a significant result in instances where the scatter ellipsoids are (1) unequally inflated, although identically oriented, (2) identically inflated but differently oriented and (3) a combination of these conditions. Equations to the major axes of a scatter ellipsoid of morphologic variables represent growth patterns in the variables. An approximate application of the asymptotic test developed by T.A. Anderson is used here to identify structure of the heterogeneity between two covariance matrices where such exists. The foregoing procedures are preliminary to a treatment of generalized distances in which the path taken by the computer program is decided by structure of the covariance matrices of the samples. Depending on the nature of the covariance matrices either the Mahalanobis' generalized distance is computed or the Anderson-Bahadur distance for heterogeneous covariance matrices. Tests of significance of the results are provided and the linear discriminant function coefficients produced as a by-product.

## INTRODUCTION

The generalized statistical distance is probably the best known and most widely used of the multivariate techniques employed in taxonomic work. It is used also widely in other connections.

Calculations performed by this program may be conceived in terms of the geometric properties of two ellipsoids, in this situation ellipsoids of scatter. For the strict application of statistical theory involved, it is necessary that the variables be distributed normally. The ellipsoid of scatter associated with a multivariate normal distribution provides a representation of the variability of the population, directly analogous to the variance of univariate statistics. In univariate statistics one will wish to make sure that in a comparison of two statistical populations on the basis of samples drawn from them the variances of the variable being analyzed are statistically identical for both populations. One also will want to know if the variables are distributed normally. The same reasoning applies in multivariate analysis. It is not difficult to test for homogeneity of univariate variances but the multivariate analog is associated with certain drawbacks and difficulties.

If the variables are distributed as a multivariate normal, the shape of a three-dimensional plot of many points will approximate a football flattened equally on two opposite sides. For a two-sample statistical comparison to be valid one would require the footballs to be of the same size and to be oriented in the same direction with respect to all axes. It is possible to employ a large-sample test to ascertain whether each principal axis of the ellipsoid of one sample is collinear with the corresponding principal axis of the other ellipsoid. In the bivariate situation there

are four categories of difference to be recognized. For three and higher dimensions the possibility of rotation about axes has to be taken into consideration.

How serious a matter is it if the covariance matrices are not equal? Experience shows that in most neontological investigations, heterogeneity in covariance matrices, although common, usually does not cause serious inaccuracies in the generalized distance. Geologic materials pose a more intricate problem. Owing to the action of geologic (nonbiologic) agencies, such as transport, considerable heterogeneity factors may be introduced into a sample. These factors may be of such an order as to cause serious inaccuracy in generalized distance computations.

Acknowledgments.—The original research work for this paper was done while R.A.R. was Visiting Research Scientist with the Kansas Geological Survey at The University of Kansas (1966-1967). The work was supported by Computing Grant 104104 at the University of Uppsala and Research Contracts 2320-26-7819G and -24-7540G.

## GEOMETRICAL INTERPRETATION

Suppose that the  $N$  vectors of a sample,  $x_1, \dots, x_N$  ( $N$  observation vectors), with mean vector  $Ex = \mu$ , and with covariance matrix  $E(x - \mu)(x - \mu)' = \Sigma$  are multivariate normally distributed and form an ellipsoid.

If one considers two  $p$ -variate normally distributed populations with the mean vectors  $\mu_1$  and  $\mu_2$ , not necessarily different (i.e.,  $\mu_1$  and  $\mu_2$  in the diagrams of Figure 1 may be coincident), and the covariance matrices,  $\Sigma_1$  and  $\Sigma_2$ , then if the test of equality of

these matrices be applied (Kullback, 1959) and it be concluded, that  $\Sigma_1$  and  $\Sigma_2$ , the following situation may prevail. In Figure 1a, the ellipsoid  $\Omega_1$  is merely a translation of ellipsoid  $\Omega_2$ . Hence,  $AB = EF$  and  $CD = GH$ , and  $AB$  is parallel to  $EF$ .

If the homogeneity test for covariance matrices leads to the conclusion, that  $\Sigma_1 \neq \Sigma_2$ , then one of the following conditions may prevail (for two dimensions). Figure 1b has  $AB$  parallel to  $EF$  and  $CD$  parallel to  $GH$ ;  $EF > AB$  and  $GH > CD$ . Hence, ellipsoid  $\Omega_2$  is a translation of ellipsoid  $\Omega_1$  with magnification.

Figure 1c has  $AB = EF$ ,  $CD = GH$ ,  $AB$  is not parallel to  $EF$  and  $CD$  is not parallel to  $GH$ . Ellipsoids  $\Omega_1$  and  $\Omega_2$  have the same shape, but are rotated in relation to each other. Figure 1d has  $AB \neq EF$ ,  $CD \neq GH$ ,  $AB$  is not parallel to  $EF$  and  $CD$  is not parallel to  $GH$ . Ellipsoids  $\Omega_1$  and  $\Omega_2$  are thus differently inflated and their axes are rotated in relation to each other.

For three or more dimensions, the situation becomes more complicated. Figure 1e illustrates the position for three dimensions. Here, two axes of ellipsoids  $\Omega_1$  and  $\Omega_2$  are parallel to each other,  $AB$  is parallel to  $EF$ , but not the remaining axes, owing to rotation about  $AB$  and  $EF$ . Hence,  $CD$  is not parallel to  $GH$  and  $IJ$  is not parallel to  $LK$ . In this example, it has been taken that the ellipsoids have the same shape, thus,  $AB = EF$ ,  $CD = GH$ , and  $IJ = KL$ . Figure 1f indicates the positions if the second axes of  $\Omega_1$  and  $\Omega_2$  are parallel and the first and third axes are rotated about the second axes. A numerical example will illustrate the foregoing. Consider the sample covariance matrices ( $N_1 = N_2 = 51$ ),  $S_1$ ,  $S_2$

$$S_1 = \begin{bmatrix} 2 & -1 \\ -1 & 1 \end{bmatrix}, S_2 = \begin{bmatrix} 2 & 1 \\ 1 & 1 \end{bmatrix}, S = S_1 + S_2 = \begin{bmatrix} 2 & 0 \\ 0 & 1 \end{bmatrix}.$$

Here,  $S$  is the pooled covariance matrix. Both of the samples have the same mean vector,  $\bar{x}_1 = \bar{x}_2 = (1, 1)'$ . The determinants of these matrices are,  $\det S_1 = 1$ ,  $\det S_2 = 1$ ,  $\det S = 2$ . The generalized test for homogeneity of covariance matrices, in the form presented in Kullback (1959, p. 317), is

$$21(H_1 : H_2(*)) = n_1 \log_e(\det S_1) + n_2 \log_e(\det S_2 / \det S), \quad (1)$$

where  $nS = n_1 S_1 + n_2 S_2$  and  $n = n_1 + n_2$  and  $n_i = N_i - 1$  ( $i = 1, 2$ ). This is approximately distributed as  $\chi^2$  with

$p(p+1)/2$  degrees of freedom, where  $p$  is the number of variables. A better approximation available, however, for only a few dimensions, is given by the  $B$ -distribution (Kullback, 1959, p. 317).

Applying (1) to the present example, one has,

$$B^2 = 50 \log_e(2/1) + 50 \log_e(2/1) = 69.315,$$

which is significant. The problem of homogeneity in covariance matrices is given detailed study in Chapter 10 of Anderson (1958).

Inasmuch as the determinants of the covariance matrices  $S_1$  and  $S_2$  are the same, both of these must have the same volume, because  $\det A = (n-1)^p \det S$  is the squared volume of a parallelotope (Anderson, 1958, p. 167). (Here,  $A$  is the matrix of sums and products,  $S$  the sample covariance matrix,  $n$  the sample size and  $p$  the number of variables.)  $S_1$  has the eigenvalues  $d_1 = 2.618$ ,  $d_2 = 0.382$  and the normalized corresponding eigenvectors are

$$b_1 = \begin{bmatrix} .8507 \\ -.5257 \end{bmatrix}, b_2 = \begin{bmatrix} .5257 \\ .8507 \end{bmatrix}.$$

Matrix  $S_2$  has the eigenvalues  $g_1 = 2.618$ ,  $g_2 = 0.382$  and the corresponding, normalized eigenvectors are

$$c_1 = \begin{bmatrix} .8507 \\ .5257 \end{bmatrix}, c_2 = \begin{bmatrix} -.5257 \\ .8507 \end{bmatrix}.$$

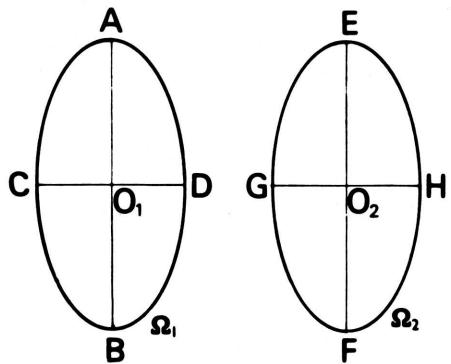
The ellipsoids have the same shape, but are oriented in relation to each other at an angle of

$$\cos \nu = 0.44733, \nu = 63^\circ 25'.$$

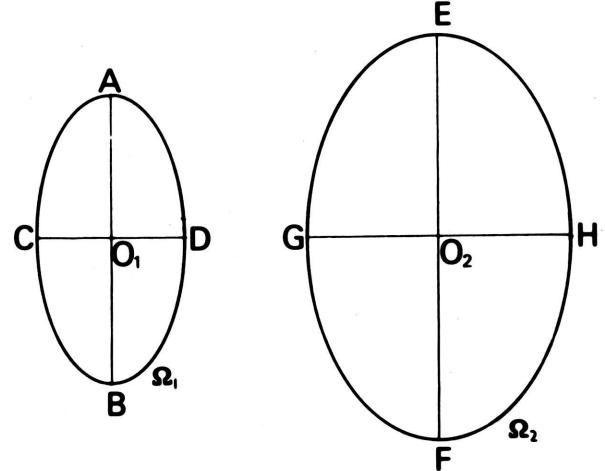
#### APPROXIMATE TEST OF ORIENTATION OF ELLIPSOIDS

In order to disclose relative heterogeneity in the orientation of the axes of two ellipsoids of scatter, if the generalized test (1) indicates  $\Sigma_1 \neq \Sigma_2$ , it is sufficient to demonstrate that the first two axes are parallel to each other in order to localize the inequality in the covariance matrices to condition (b) of Figure 1. This demonstrates that the heterogeneity is due solely to a greater degree of scatter in the points of the observational vectors forming the ellipsoids.

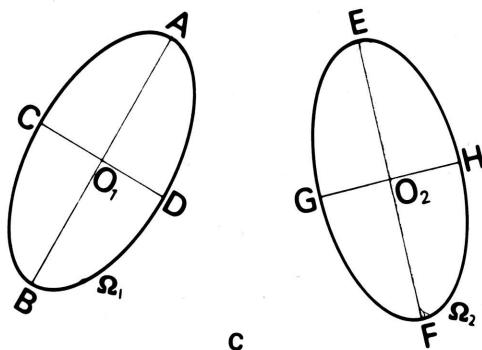
Some interest may attach to comparing the relative differences in the inflations of ellipsoids  $\Omega_1$  and  $\Omega_2$ . This may be gauged approximately by comparing the eigenvalues of matrices  $\Sigma_1$  and  $\Sigma_2$ . Let



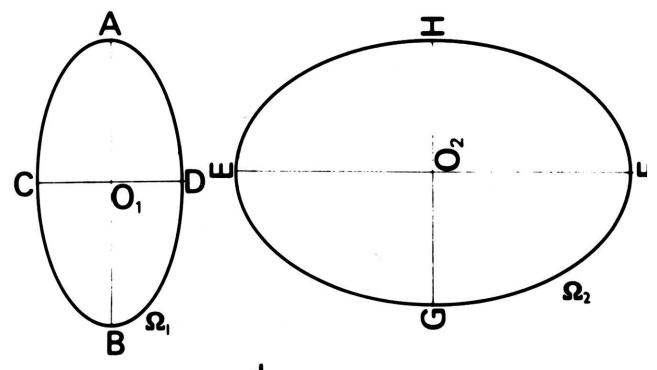
a



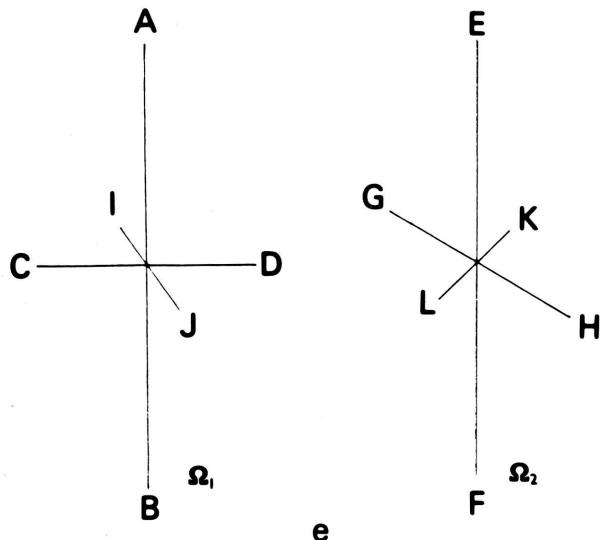
b



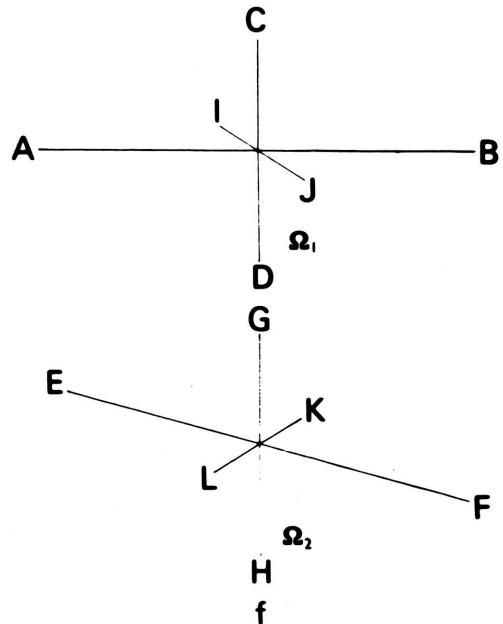
c



d



e



f

Figure 1. - a, Ellipsoids in translation; b, ellipsoids in translation and one with magnification; c, ellipsoids in rotation; d, ellipsoids in rotation and one with magnification; e, position for three dimensions; f, three-dimensional position with two axes rotated.

$\delta_1 > \dots > \delta_p > 0$  be the  $p$  eigenvalues of the positive definite matrix  $\Sigma_1$ ; then  $|\Sigma_1 - \delta I| = 0$ , and  $\gamma_1, \dots, \gamma_p$  are the corresponding normalized eigenvectors which satisfy  $\Sigma_1 \gamma_i$  and  $\delta_i \gamma_i$  and  $\gamma_i \gamma_i = 1$ . Likewise let  $\lambda_1 > \dots > \lambda_p > 0$  be the  $p$  eigenvalues of the positive definite matrix  $\Sigma_2$ ;  $|\Sigma_2 - \lambda I| = 0$  and  $\beta_1, \dots, \beta_p$  the corresponding normalized eigenvectors which satisfy  $\Sigma_2 \beta_i = \lambda_i \beta_i$  and  $\beta_i \beta_i = 1$ . If the  $\delta_i$  and  $\lambda_i$  are different,  $\gamma_i \gamma_i = 0$  and  $\beta_i \beta_i = 0$ . The length of the  $i$ -th principal axis of the ellipsoid of scatter is related to the magnitude of the corresponding eigenvalue, in the sense that it constitutes the variance of the  $i$ -th principal component.

For the first eigenvalue and eigenvectors of covariance matrix  $\Sigma_1$  we have that

$$\delta_1 \gamma_1 \Sigma_1^{-1} \gamma_1 = \delta_1 \delta_1^{-1} = 1,$$

and,

$$(1/\delta_1) \gamma_1 \Sigma_1 \gamma_1 = (1/\delta_1) \delta_1 = 1,$$

and similarly for  $\Sigma_2$ . If, say, the first eigenvector of  $\Sigma_2$ ,  $\beta_1$ , has the same direction cosines as the first eigenvector of  $\Sigma_1$ ,  $\gamma_1$ , then  $\delta_1 \beta_1 \Sigma_1^{-1} \beta_1 = 1$  and  $(1/\delta_1) \beta_1 \Sigma_1 \beta_1 = 1$ . Anderson (1963, p. 144) has put forward a procedure for testing the null hypothesis, that a given eigenvector (principal component) is a specified vector. The procedure is based on the limiting normal distribution possessed by  $\sqrt{n}(c_1 - \gamma_1)$ , where  $\gamma_1$  is the first eigenvector of a population covariance matrix and  $c_1$  a sample estimate thereof, based on a sample of size  $(n+1)$ . As an approximate test of the hypothesis that a given eigenvector is the  $i$ -th eigenvector of a large sample estimate of a covariance matrix, it is suggested, that the following version of Anderson's test criterion might be applicable

$$n(d_i b_i' S_1^{-1} b_i + (1/d_i) b_i' S_1 b_i - 2). \quad (2)$$

Here,  $(n+1)$  is the sample size of  $S_1$ , the estimate of  $\Sigma_1$ ,  $d_i$  is the sample estimate of  $\delta_i$  and  $b_i$  that of  $\beta_i$ , the latter being based on a large sample. If  $\Sigma_1 = \Sigma_2$ , then  $\gamma_i = \beta_i$ . This criterion has, of course, more variables than the original form given by Anderson and must, therefore, be more susceptible to fluctuations.

Test (1) may then be computed. The quadratic forms of (2) may be calculated separately and the test worked out as a subsequent step. The first test criterion is asymptotically distributed as  $\chi^2$  with  $p$

$(p+1)/2$  degrees of freedom, where  $p$  is the number of variables. The Anderson test criterion is distributed as  $\chi^2$  with  $(p-1)$  degrees of freedom. In judging the significance of a result, it seems advisable to take the 1% level as the highest level of risk, rather than the customary 5%.

## GENERALIZED DISTANCE CALCULATIONS

The Mahalanobis' generalized statistical distance squared is yielded by the relationship

$$D^2 = d' S^{-1} d, \quad (3)$$

where  $d$  is the vector of differences between the sample mean vectors of the two samples and  $S$  is the pooled covariance matrix for the two samples. The statistical significance of this result is assessed by means of the Hotelling  $T^2$ , which is related to  $D^2$  as follows

$$T^2 = D^2 \frac{N_1 N_2}{N_1 + N_2}, \quad (4)$$

where  $N_1$  and  $N_2$  are the sizes of the two samples involved. The final step in finding the significance of the generalized distance is the calculation of the variance ratio ( $F$ ) by means of the following formula

$$F = T^2 \frac{(N_1 + N_2 - p - 1)}{p(N_1 + N_2 - 2)}. \quad (5)$$

Here  $p$  represents the number of variables.

Anderson and Bahadur (1962) considered a generalized statistical distance for unequal covariance matrices of the following form

$$D_H = \frac{2b'd}{(b'S_1 b) + (b'S_2 b)}. \quad (6)$$

Here,  $d$  is again the difference mean vector,  $b$  is an analog of the vector of coefficients of the Fisherian discriminant function and  $S_1$  and  $S_2$  are the respective samples covariance matrices. If the covariance matrices are equal this formula reduces to formula (3). Formula (6) is not connected with a test of significance. In order to obtain a significance test based on  $T^2$  where the covariance matrices are unequal one may proceed by means of a distance method proposed by Reymert (1962).

## DESCRIPTION OF PROGRAM

### Operating Instructions

The following INPUT cards are required.

Card 1: The number of JOBS occurring in the current sequence in FORMAT (3X, 12).

Card 2: The following information is punched on this card in FORMAT (25I3), in the

- order given.
- Column 1-3. M = number of variables.  
 4-6. NA = size of the larger sample.  
 7-9. NB = size of the smaller sample.  
 10-12. MATRIX = 0 if observation vectors are read in and = 1 for data in matrix form.  
 13-15. L = 1 for cross-products matrices and = 0 for covariance.  
 16-18. LOGDEC = 1 if the data are transformed to the base ten logarithms and = 0 for raw data.  
 19-21. LDISCR = 1 if the Mahalanobis' generalized distance is computed in addition to the heterogeneous distance if this branch has been chosen by the course of the program, otherwise put = 0.  
 22-24. INVST (1) } intervals for histogram subroutines  
 25-27. INVST (2) }  
 28-30. INVST (3) }  
 31-33. KREYD. Supply here the number of times program REYDST is to operate.

Card 3: The TITLE card.

Card 4: The FORMAT card for reading the data.

Card 5 and subsequent cards: either the data  $x(I)$  in matrix form in accordance with the format given by Card 4 or the data as vectors of observations. Sample A is always to be located before sample B, the smaller sample. Thus, for the matrix option, matrices A (I, J) and B (I, J) and vectors SA (I) and SB (I) are read in at this point. The larger sample should be greater than 100 observations for the theory used in this program to have full validity.

### Mainline Program ORNTDIST

Calculations either are made directly on the covariance matrices (or cross-products matrices) of samples A and B and the corresponding mean vectors or such are computed from the observations. The main program produces the basic information required for the statistical tests and calculations. In addition to the covariance matrices and mean vectors the program computes the eigenvalues and eigenvectors of matrices A and B as well as those of matrix C, obtained by pooling A and B.

### Subroutines BAHADU and HETDST

The subroutines perform the heterogeneous distance calculations of Anderson and Bahadur (1962). The value of the expression

$$\frac{2b'd}{\sqrt{(b'S_1b)} + \sqrt{(b'S_2b)}}$$

is found, where d is the difference vector between sample means,  $S_1$  and  $S_2$  are the sample covariance ma-

trices and b is a vector estimated iteratively in the two subroutines by trial and error solutions of

$$b = \{tS_1 + (1-t)S_2\}^{-1}d, \quad (7)$$

t being a scalar term located between zero and unity. The value of t, initially guessed, is improved iteratively from

$$b' \{t^2 S_1 - (1-t)^2 S_2\} b = 0. \quad (8)$$

Convergence is checked for in HETDST. For each iteration the output from BAHADU consists of the heterogeneous generalized distance, its square; t and the iteration factors are given from HETDST. The analog of the coefficients of the discriminant function SB(I) also are printed out.

### Subroutine DISTFN

The subroutine computes the Mahalanobis' generalized distance for homogeneous covariance matrices. The output consists of the discriminant function coefficients, the generalized distance, its square, Hotelling's  $T^2$ , the corresponding variance ratio and degrees of freedom for it. The calculations are based essentially on the expansion of a quadratic form.

### Subroutine HAFCOV

The subroutine computes a heterogeneous generalized distance by averaging the covariance matrices  $S_1$  and  $S_2$ . The calculations then proceed as in DISTFN.

The output comprises a generalized distance and its square. HAFCOV provides a valuable check on the calculations in other parts of the program but it is not a vital section and may be excluded from the program by punching C in column 1 of CALL HAFCOV in subroutine HETDST.

### Subroutine ORIENT

The subroutine is called from the main program. It operates on eigenvectors of the comparison matrix, in conjunction with the reference matrix, to test the significance in differences in orientation of the scatter ellipsoids. The computations are based on formula (2). The subroutine does not treat eigenvectors whose eigenvalues are less than one percent of the trace.

### Subroutine REYDST

The role of the subroutine is to calculate a formal test of significance for the differences in means if the covariance matrices are not equal. It is based on an iterative randomization procedure and is operated in conjunction with subroutines RANKRT and CLEAR.

REYDST is called from HETDST. The details of the calculations are the same as for DISTFN. The subroutine is erratic if deviations from multivariate normality occur and tends to overestimate the generalized distance. This becomes particularly noticeable for many dimensions. The calculations are time-consuming owing to the randomization procedure. For four variables and 30 randomizations the entire program took 5.36 minutes of which about 85% was taken by REYDST. Clearly, the user of the program may consider this to be too uneconomical to make the significance test worthwhile. The instruction KREYD if put = 0 causes REYDST to be omitted. Otherwise the number of iterations required is specified here.

#### Subroutine RANKRT

This selects cards randomly from the stored decks of data cards. The use of RANKRT requires two random access (RA) control cards on the CDC 3600 - e.g.

{7  
{9EQUIP, 10 = RA(10000). CLEAR must be called to initiate the use of RANKRT.

#### Subroutine HOUSER

This is a standard subroutine for computing eigenvalues and eigenvectors of a square symmetric matrix by the Householder method. The matrix is reduced to tridiagonal form, the eigenvalues of the tridiagonal form are computed then by the Sturm sequence process and finally the eigenvectors are obtained by the Wilkinson inverse iteration method.

#### Subroutine PUTMAT

This subroutine writes a matrix on a scratch unit as one logical record. The matrices A, B, and C are stored in this manner.

#### Subroutine GETMAT

This reads one of the matrices A, B, or C from the scratch unit generated by PUTMAT.

#### Subroutine MATOUT

This prints out a matrix on the printer unit. Columns

and rows are marked.

#### Subroutine JZUP2

This subroutine carries out a substitution of the original measurements, or the logarithmically transformed measurements, into the vector of discriminant function coefficients. The values are plotted as a histogram by PLTHS1.

#### Subroutine MATINV

This is a standard subroutine for carrying out the inversion of square, symmetric matrices.

#### Subroutine DISTRB

This calculates the coefficients of skewness and kurtosis for each of the variables for both samples. The large-sample standard deviations of each coefficient are found and an approximate t-value calculated. The foregoing results as well as the means and standard deviations for each variable are printed out.

#### Subroutine PLTHS1

This is a standard subroutine of the Computation Center of The University of Kansas for printing histograms. It is called from the subroutines DISTFN, and HETDST through JZUP2. The resulting histogram of substituted values gives a valuable pictorial representation of the efficiency of the discriminant function in separating the two samples.

#### NOTE ON PROBLEM EXAMPLE

The example used to illustrate the program consists of measurements on male and female grasshoppers from Gotland (Omocestus haemorrhoidalis L.). The variables are  $X_1$  = length of hind femur,  $X_2$  = pronotal length,  $X_3$  = elytron length and  $X_4$  = the least width between the ridges of the pronotus. The example used here is an entomological one, but any type of normally distributed multivariate data are amenable to distance analysis by this program, and it may equally as well be applied to sedimentologic, anthropometric, geochemical and sociologic studies just to mention a few possibilities (Reyment, 1969).

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APPENDIX A. -Program Listing.

LISTABLE OUTPUT FROM SEQUENCE 003027  
 UDAC DRUM SCOPE 2,020, TWO DRUMS  
 SEQUENCE NUMBER 003027 STARTED AT TIME 063358 DATED 28/03-69  
 JOB,104104,REYMENT,3, BUD PALEONTOLOGEN  
 DEMAND,23000  
 EQUIP,10=RA(25000)  
 EQUIP,11=RA(25000)  
 EQUIP,26=(104104 ORNTDIST,,,999)  
 FTN,L,X=26,\*

PROGRAM ORNTDIST

LIST OF SUBROUTINES		
HOUSER	= COMPUTES EIGENVALUES AND EIGENVECTORS	
ORIENT	= COMPARES ORIENTATIONS OF EIGENVECTORS	
HETDST	= COMPUTES ANDERSON/BAHADUR DISTANCE	
BAHADU	= PERFORMS STEPS IN THE ITERATIVE CALCULATION OF HETDST	
REYDST	= COMPUTES SIGNIFICANCE OF A HETEROGENEOUS DISTANCE	
HAFCOV	= GIVES A HETEROGENEOUS DISTANCE BY AVERAGING COVARIANCES	
DISTFN	= FINDS THE MATANOLOBIS DISTANCE	
RANKRT	MATINV	JZUP2
PUTMAT	GETMAT	MATOUT
PLTHS1	= PLOTS HISTOGRAMS	
DISTRB	= TESTS THE DISTRIBUTIONS OF EACH OF THE VARIABLES	

EXPLANATION OF CONTROL CARDS

CARD 1      JOBS      = NUMBER OF JOBS BEING RUN IN SEQUENCE  
 CARD 2  
 COL 1-3      M = NUMBER OF VARIABLES (MAXIMUM = 45)  
 COL 4-6      NA = IS SIZE OF THE LARGER SAMPLE (MAXIMUM=500)  
 COL 7-9      NB = SIZE OF SMALLER SAMPLE (MAXIMUM=500)  
 COL 10-12     MATRIX = 0 FOR SERIES OF MEASUREMENTS AND 1 FOR  
 DATA READ IN IN MATRIX FORM ,I,E, PARTIALLY PROCESSED)  
 COL 13-15     L = 1 FOR INPUT DATA AS CROSS-PRODUCTS MATRICES  
 L = 0 FOR COVARIANCE MATRICES AND SERIES OF  
 MEASUREMENTS  
 COL 16-18     LOGDEC = 1 IF LOGARITHMIC TRANSFORMATION OF RAW DATA  
 REQUIRED AND 0 IF THE DATA ARE NOT TO BE TRANSFORMED  
 COL 19-21     LDISC = 1 IF THE NORMAL FORM OF THE GENERALIZED  
 DISTANCE IS TO BE COMPUTED IN ADDITION TO THE  
 DECISION MADE BY THE PROGRAM = THIS IS A USEFUL  
 CONTROL STEP  
 COL 22-24, 25-27, 28-30    INVST = NUMBER OF INTERVALS FOR EACH  
 HISTOGRAM, IF NO SELECTION ALTERNATIVELY ZEROES  
 PROGRAM PICKS INTERVAL OF TEN  
 COL 31-33     KREYD = NUMBER OF TIMES IN THE REYDST=LOOR  
 CARD 3      THE TITLE CARD WITH THE HEADING IN 72 SPACES  
 CARD 4      THE VARIABLE FORMAT CARD FOR THE DATA  
 CARD 5      THE DATA IN ONE OF THE PRESELECTED FORMS  
 THE LARGER SAMPLE MUST ALWAYS BE READ FIRST  
 DIMENSION E(45,45),F(45,45),AEGVAL(45),EEGVAL(45),CEGVAL(45),  
 1 TDATA(30),FMT(12),TITLE(12)  
 COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),  
 1 T(45),X(45),Z(45),M,NA,NE,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,  
 2 MATPOS,KREYD

```

COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMAEC,MTARA,MTBRA
DATA(TDATA# 0.0,9.21,11.34,13.28,15.09,16.781,18.48,20.09,21.67,23,
121,24.72,26,22,27,69,29,14,30,58,32.00,33,41,34,81,36,19,37,57,38,
293,40.29,41,64,42,98,44,31,45,64,46,96,48,28,49,59,50,89)
C=F IV- DATA TDATA/.0,9.21,11.34,13.28,15.09,16.81,18.48,20.09,21.67,23,
121,24.72,26,22,27,69,29,14,30,58,32.00,33,41,34,81,36,19,37,57,38,
293,40.29,41,64,42,98,44,31,45,64,46,96,48,28,49,59,50,89/
C
C
C----- LIN IS THE STANDARD INPUT UNIT (CARD READER)
LINE=60
C----- LUT IS THE STANDARD OUTPUT UNIT (PRINTER)
LUTE=61
C----- MTA,MTB,MTMAEC,MTARA AND MTBRA ARE SCRATCH UNITS
C----- ***** SCRATCH TAPE USAGE *****
C----- TAPE MTA CONTAINS SAMPLE A
MTA=51
C----- TAPE MTB CONTAINS SAMPLE B
MTB=52
C----- MTMABC CONTAINS MATRICES A,B AND C IN THIS ORDER,ONE PER RECORD
MTMABC=50
C----- RANDOM ACCES UNITS ARE USED ONLY IN KEYDST AND RANKRT
C----- MTARA IS A RANDOM ACCESS UNIT WHERE SAMPLE A IS STORED
MTARA=10
C----- MTBRA IS A RANDOM ACCESS UNIT WHERE SAMPLE B IS STORED
MTBRA=11
READ (LIN,380) JOBS
KLOBS = 0
MW=45
C----- HERE BEGINS A NEW CYCLE
1111 CONTINUE
REWIND MTA
REWIND MTB
REWIND MTMABC
READ (LIN,3) M,NA,NB,MATRIX,L,LOGDEC,LDISCR,INVST,KREYB
IF (M .LE. 0) CALL EXIT
IF (M .LE. MW .AND. NA .LE. 500 .AND. NE .LE. 500) GO TO 33
WRITE (LUT,132) M,NA,NE
33 DO 10101 IX = 1,3
10101 IF(INVS1(IX).LE.0)INVST(IX) = 10
DO 1620 I=1,M
SA(I)=0,0
SB(I)=0,0
DO 1620 J=1,M
A(I,J)=0,0
B(I,J)=0,0
C(I,J)=0,0
1620 CONTINUE
WRITE (LUT,207)
IF (LOGDEC ,EQ, 1) WRITE (LUT,490)
EM=M
ENA=NA
ENB=NB
READ (LIN,1) TITLE
WRITE (LUT,333) TITLE
READ (LIN,1) FMT
WRITE (LUT,400) FMT
IF(MATRIX.GT.0) GO TO 5000
WRITE (LUT,710) NA,NB
WRITE (LUT,5100)

```

```

NABENA
DO 6009 IREP=1,2
C===== WHEN IREP .EQ. 1 STORE SAMPLE A ON MTA
C===== WHEN IREP .EQ. 2 STORE SAMPLE B ON MTB
LUN#MTA
IF (IREP .EQ. 2) LUN#MTB
DO 6008 I=1,NAB
READ (LIN,FMT) (X(IX),IX=1,M)
WRITE (LUT,5020) (X(IX),IX=1,M)
IF (LOGUEC .EQ. 0) GO TO 5051
DO 5050 IX=1,M
5050 X(IX)= ALOG10(X(IX))
5051 WRITE (LUN) (X(IX),IX=1,M)
625 CONTINUE
627 IF (IREP,EQ.2) GO TO 635
DO 626 I=1,M
SA(I)=SA(I)+X(I)
DO 626 J=1,M
626 A(I,J)=A(I,J)+X(I)*X(J)
GO TO 6008
635 DO 628 I=1,M
SB(I)=SB(I)+X(I)
DO 628 J=1,M
628 B(I,J)=B(I,J)+X(I)*X(J)
6008 CONTINUE
IF (IREP,EQ.2) GO TO 6009
NAB#NB
WRITE (LUT,5105)
6009 CONTINUE
655 DO 660 I=1,M
DO 660 J=1,M
A(I,J)=(A(I,J)-SA(I)*SA(J)/ENA)/(ENA-1.)
660 B(I,J)=(B(I,J)-SB(I)*SB(J)/ENB)/(ENB-1.)
DO 199 IX=1,M
SA(IX)=SA(IX)/ENA
199 SB(IX)=SB(IX)/ENB
CALL DISTRB
GO TO 1117
5000 WRITE (LUT,5100)
DO 10 I=1,M
10 READ (LIN,FMT) (A(I,J),J=1,M)
CALL MATOUT (A)
WRITE (LUT,5105)
DO 30 I=1,M
30 READ (LIN,FMT) (B(I,J),J=1,M)
CALL MATOUT (B)
READ (LIN,FMT) (SA(IX),IX=1,M)
READ (LIN,FMT) (SB(IX),IX=1,M)
IF (L,LE,0) GO TO 1117
1116 DO 1121 I=1,M
DO 1121 J=1,M
1121 A(I,J)=A(I,J)/(ENA-1.)
DO 1122 I=1,M
DO 1122 J=1,M
1122 B(I,J)=B(I,J)/(ENB-1.)
1117 DO 40 I=1,M
DO 40 J=1,M
40 C(I,J)=((ENA-1.)*A(I,J)*(ENB-1.)*B(I,J))/(ENA*ENB-2.)
WRITE (LUT,1120)
CALL MATOUT (A)
C      MATRIX A IS REFERENCE

```

```

      WRITE (LUT,1119)
      CALL MATOUT (B)
      WRITE (LUT,1118)
      CALL MATOUT (C)
      CALL PUTMAT(A)
      DO 5432 IX=1,M
5432 SC(IX)=SB(IX)-SA(IX)
      WRITE (LUT,3333)
      WRITE (LUT,3333)
      CALL Houser (M,M,A,AEGVAL,1.,E)
      WRITE (LUT,716)
      WRITE (LUT,1492)
      WRITE (LUT,20) (AEGVAL(IX),IX=1,M)
      WRITE (LUT,1493)
      CALL MATOUT (E)
      DETAE=0.
      DO 80 IX=1,M
      IF (AEGVAL(IX) .LE. 1.E-09) GO TO 561
80 DETAE=DETA+ ALOG(AEGVAL(IX))
      CALL PUTMAT(B)
C     MATRIX B IS STORED
      WRITE (LUT,3333)
      CALL Houser (M,M,B,BEGVAL,1.,F)
      WRITE (LUT,1494)
      WRITE (LUT,20) (BEGVAL(IX),IX=1,M)
      WRITE (LUT,1495)
      CALL MATOUT (F)
      DETBF=0.
      DO 100 IX=1,M
      IF (BEGVAL(IX) .LE. 1.E-09) GO TO 561
100 DETBF=DETB+ ALOG(BEGVAL(IX))
      CALL PUTMAT(C)
      MATPOS=3
      WRITE (LUT,3333)
      WRITE (LUT,3333)
      CALL Houser (M,M,C,CEGVAL,1.,F)
      WRITE (LUT,1497)
      WRITE (LUT,20) (CEGVAL(IX),IX=1,M)
      WRITE (LUT,1498)
      CALL MATOUT (F)
      DETC=0.
      DO 120 IX=1,M
      IF (CEGVAL(IX) .LE. 1.E-09) GO TO 561
120 DETC=DETC+ ALOG(CEGVAL(IX))
      WRITE (LUT,90)
      WRITE (LUT,110) DETA,DETE,DETC
      BSQ=(ENA-1.)*(DETC-DETA)*(ENB-1.)*(DETC-DETB)
      BETASQ=((2.*FLOAT(M)**3+3.*FLOAT(M)**2-FLOAT(M))/12.)*
1((1./(ENA-1.))+1./((ENB-1.)*1.)/(ENA+ENB-2.))
      WRITE (LUT,716)
      WRITE (LUT,140) BSQ,BETASQ
      KT = M*(M+1)/2
      WRITE (LUT,151) KT
      TRACE = 0.
      DO 170 IX=1,M
170 TRACE=TRACE+BEGVAL(IX)
      WRITE (LUT,180) TRACE
      DO 200 IX=1,M
200 Z(IX)=(BEGVAL(IX)*100.)/TRACE
      WRITE (LUT,220) (Z(IX),IX=1,M)

```

```

CALL GETMAT (E,2)
DO 500 I=1,M
DO 500 J=1,M
500 A(I,J)=B(I,J)
C      MATRIX A IS FOR INVERSION
CALL MATINV (A,M,CHGVAL,0,DETERM,MW)
KOUNT = 1
WRITE (LUT,160)
DO 230 KL = 1,M
IF(Z(KL).LE.1.) GO TO 10000
CALL ORIENT (E,B,M,X,A,KOUNT,ENB,BEGVAL)
KOUNT = KOUNT + 1
230 CONTINUE
10000 CALL GETMAT(A,1)
CALL GETMAT(B,2)
IF(KT.GT.30) GO TO 5551
TEST = TDATA(KT)
WRITE (LUT,10598) TEST
GO TO 5552
5551 TEST=0.5*(2.*((ENA-ENB-1.)*0.96)
WRITE (LUT,425) TEST
5552 IF(TEST.GT.BSQ) CALL DISTFN
IF(TEST.LE.BSQ) CALL METDIST
GO TO 1000
561 WRITE (LUT,562)
1000 KLOBS = KLOBS + 1
IF(JCBS-KLOBS) 365,365,1111
365 WRITE (LUT,360)
CALL EXIT
C
C
1 FORMAT(12A6)
3 FORMAT (25I3)
20 FORMAT(1X,11F12.5)
90 FORMAT(27H0      LOG DETERMINANTS /41H   REFERENCE      COMPARISON
       1N      POOLED    )
110 FORMAT (2X,F7.3,6X,F7.3,9X,F7.3)
133 FORMAT (30H0NUMBER OF VARIABLES TOO HIGH= ,I4,5H, NAM,14:4H,NB,I4)
140 FORMAT(17H      B SQUARE = ,F14.7/20H      BETA SQUARE = ,F14.7)
151 FORMAT(27H      DEGREES OF FREEDOM = I3)
160 FORMAT(31H1      ORIENTATION OF ELLIPSOIDS      )
180 FORMAT(13H0      TRACE B = F18.7)
207 FORMAT(50H1      ANALYSIS OF HOMOGENEITY OF COVARIANCE MATRICES)
220 FORMAT(23H0      PERCENTAGES FOR B(I)/(X,8F14.7))
333 FORMAT(10X12A6)
360 FORMAT(18H1      CYCLE COMPLETED/12H0  RETURNING)
380 FORMAT(3X,I2)
400 FORMAT (38H0DATA CARDS READ WITH VARIABLE FORMAT ,12A6)
425 FORMAT(23H0      THE FORKING VALUE = F14.7)
490 FORMAT(29H0      DATA LOGARITHM TRANSFORMED)
562 FORMAT(50H0NEGATIVE EIGENVALUE, PROCEEDING TO NEXT DATA SET )
710 FORMAT(48H0      REFERENCE SAMPLE SIZE BASED ON POPULATION OF ,I3/29H0
       1      COMPARISON SAMPLE SIZE IS ,I3)
716 FORMAT(1H0)
1118 FORMAT(20H0 ROW POOLED MATRIX)
1119 FORMAT(47H0 ROW COVARIANCE MATRIX 2 ( COMPARISON MATRIX )  )
1120 FORMAT(47H0 ROW COVARIANCE MATRIX 1 ( REFERENCE MATRIX )  )
1492 FORMAT(35H0 EIGENVALUES FOR REFERENCE MATRIX)
1493 FORMAT(36H0 EIGENVECTORS FOR REFERENCE MATRIX)
1494 FORMAT(36H0 EIGENVALUES FOR COMPARISON MATRIX)

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1495 FORMAT(37H0 EIGENVECTORS FOR COMPARISON MATRIX)
1497 FORMAT(32H0 EIGENVALUES FOR POOLED MATRIX)
1498 FORMAT(33H0 EIGENVECTORS FOR POOLED MATRIX)
3333 FORMAT(6H0****)
5020 FORMAT(1X10F13.7)
5100 FORMAT(23H0 FIRST DATA INPUT SET)
5105 FORMAT(24H0 SECOND DATA INPUT SET)
10598 FORMAT(10H0 TEST = F14.7)
END

```

SUBROUTINE GETMAT (STMAT,MATNR)

```

C----- GETMAT READS THE MATNR.TH MATRIX FROM LOGICAL UNIT MTMABC TO
C----- THE PARAMETER STMAT.
C----- MTMABC CONTAINS 3 RECORDS WHICH CONTAINS RESP MATRIX A,
C----- MATRIX B AND MATRIX C,
C----- IN THE VARIABLE MATPOS IS INDICATED WHICH MATRIX WAS READ LAST.
C----- GETMAT MUST NOT BE CALLED UNTIL PUTMAT HAS WRITTEN ALL 3
C----- MATRICES E.G. PUTMAT HAS BEEN CALLED THREE TIMES.
DIMENSION STMAT(45,45)
COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
1 T(45),X(45),Z(45),M,NA,NB,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
2 MATPOS,KREYD
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
L=MATNR
C----- J IS THE NUMBER OF RECORDS TO SKIP FORWARD (IF J .GT. 0) OR
C----- BACKWARD (IF J .LT. 0)
J=L-1-MATPOS
C----- IF J .EQ. 0 MTMABC IS ALREADY POSITIONED.
IF (J) 10,20,30
10 J=-J
C----- BACK OVER J RECORDS
DO 11 J1=1,J
11 BACKSPACE MTMABC
GO TO 20
C----- SKIP OVER J RECORDS
30 DO 31 J1=1,J
31 READ (MTMABC) J2
C----- READ THE MATRIX
20 READ (MTMABC) ((STMAT(I,JC),JC=1,M),I=1,M)
C----- UPDATE MATPOS
MATPOS=L
RETURN
END

```

SUBROUTINE PUTMAT(STMAT)

```

C----- PUTMAT WRITES THE MATRIX STMAT ON LOGICAL UNIT MTMABC.
C----- PUTMAT EXPECTS THAT MTMABC IS REWOUND BEFORE FIRST MATRIX IS
C----- STORED.
C----- FIRST CALL IS FOR STORING MATRIX A, SECOND FOR MATRIX B AND
C----- THIRD FOR MATRIX C.
C----- EACH MATRIX WILL OCCUPY ONE RECORD.
DIMENSION STMAT(45,45)
COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
1 T(45),X(45),Z(45),M,NA,NE,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
2 MATPOS,KREYD
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
WRITE (MTMABC)((STMAT(I,J),J=1,M),I=1,M)
RETURN
END

```

```

SUBROUTINE Houser (N,NEV,A,C,EV,VEC,V)
C----- COMPUTES EIGENVALUES-EIGENVECTORS---SYMMETRIC MATRIX
DIMENSION A(45),B(50),C(45,45),EV(45),P(45),TA(50),TH(50),W(50),
1 Y(45),V(45,45)
NN=N-1
DO 16 I=1,NN
SUM1=0
B(I)=0
JI=I+1
DO 14 J=JI,N
14 SUM1 = SUM1 + (C(I,J)*C(I,J))
S=SQRT(SUM1)
IF(S)15,16,15
15 SGN=SIGN(1.,C(I,I+1))
TEMP=SGN*(C(I,I+1))
W(I+1)=SQRT(.5*(1.0+(TEMP/S)))
C(I,I+1)=W(I+1)
II=I+2
IF(II-N)250,250,260
250 TEMP=SGN/(2.*W(I+1)*S)
DO 20 J=II,N
W(J)=TEMP*C(I,J)
20 C(I,J)=W(J)
260 B(I)=-SGN*S
DO 22 L=JI,N
SUM2=0.0
DO 21 M=JI,N
21 SUM2 = SUM2 + (C(L,M) * W(M))
22 P(L)=SUM2
XKAP=0.0
DO 23 K=JI,N
23 XKAP=XKAP + (W(K) * P(K))
DO 24 L=JI,N
24 P(L)=P(L)-(XKAP * W(L))
DO 26 J=JI,N
DO 25 K=J,N
25 C(J,K)=C(K,J)-(2.0*((P(J) * W(K)) + (P(K) * W(J))))
C(K,J)=C(J,K)
26 CONTINUE
16 CONTINUE
17 DO 18 K=1,N
18 A(K)=C(K,K)
B(N-1)=-B(N-1)
B(N)=0.0
DO 500 I=1,N
500 W(I)=H(I)
C----- STURM METHOD (MODIFIED FROM SAND HCUSES.R)
29 U=ABS(A(1))+ABS(B(1))
DO 30 I=2,N
BD=ABS(A(I))+ABS(B(I))+ABS(B(I-1))
30 IF (BD .GT. U) U=BD
U2=U**2
DO 32 I=1,N
B(I)=B(I)**2/U2
A(I)=A(I)/U
EV(I)=-1.
32 CONTINUE
BD=U
U=1.
DO 1500 K=1,NEV

```

```

38 EL=EV(K)
ELAM=.5*(U+EL)
IF (ABS(ELAM-U) .LT. 1.E-20) GO TO 34
33 IF (ABS(ELAM-EL) .LT. 1.E-20) GO TO 34
35 P0=1.
P1=A(1)-ELAM
ZNSIG2=1.
B2=0,
B1=B(1)
IF(P1)1051,1052,1052
1051 ZNSIG1=-1,
IAG=0
GO TO 1053
1052 ZNSIG1=1,
IAG=1
1053 DO 100 I=2,N
ALPH=A(I)-ELAM
IF(B1)115,111,115
111 P2=ZNSIG1*ALPH
GO TO 114
115 IF(B2)116,117,116
116 IF (ABS(P1)+ABS(P0) .GE. 1.E-20) GO TO 152
151 P1=1,E20*P1
P0=1,E20*P0
152 P2=ALPH*P1-B1*P0
GO TO 114
117 P2=ALPH*P1-ZNSIG2*B1
114 P0=P1
B2=B1
B1=B(I)
P1=P2
ZNSIG2=ZNSIG1
IF(P2) 121,125,122
121 ZNSIG1=-1,
GO TO 123
122 ZNSIG1=1.
123 IF (ZNSIG1*ZNSIG2 ,LE. 0) GO TO 100
125 IAG=IAG+1
100 CONTINUE
IF (IAG .GE. K) GO TO 40
42 U=ELAM
GO TO 38
40 M=K+1
DO 41 MM=M,IAG
41 EV(MM)=ELAM
EL=ELAM
GO TO 38
34 EV(K)=ELAM
1500 CONTINUE
DO 1060 I=1,N
A(I)=A(I)*BD
EV(I)=EV(I)*BD
1060 CONTINUE
43 IF(VEC) 44,700,44
44 L = NEV - 1
DO 502 K = 1,L
IF (EV(K) .GT. EV(K+1)) GO TO 502
501 EV(K+1)=.999999999*EV(K)
502 CONTINUE
DO 600 I = 1,NEV
II = 0

```

```

DO 503 J=1,N
503 Y(J)=1,
601 DO 504 K=1,N
P(K) = 0.0
TB(K) = W(K)
504 TA(K) = A(K) - EV(I)
L = N-1
DO 505 J = 1,L
IF (ABS(TA(J))-ABS(W(J))) 507,506,506
506 F = W(J) / TA(J)
GO TO 509
507 F = TA(J)/W(J)
TA(J) = W(J)
T = TA(J+1)
TA(J+1) = TB(J)
TB(J) = T
P(J) = TB(J+1)
TB(J+1)=0,0
IF (II-1) 509,508,509
508 T = Y(J)
Y(J) = Y(J+1)
Y(J+1) = T
509 TB(J+1) = TB(J+1)-F*P(J)
TA(J+1) = TA(J+1)-F*TB(J)
IF (II-1) 505,510,505
510 Y(J+1) = Y(J+1) - F*Y(J)
505 CONTINUE
IF (TA(N)) 511,512,511
512 TA(N) = 10.E-30
511 IF (TA(N-1)) 513,514,513
514 TA(N-1) = 10.E-30
513 Y(N) = Y(N)/TA(N)
Y(N-1) = (Y(N-1)-Y(N)*TB(N-1))/TA(N-1)
L = N-2
DO 515 J = 1,L
K = N-J+1
IF (TA(K)) 515,516,515
516 TA(K) = 10.E-30
515 Y(K) = (Y(K) - Y(K+1)*TB(K)-Y(K+2)*P(K))/TA(K)
IF (II) 517,518,517
518 II = 1
GO TO 601
517 DO 521 J=1,L
T = 0.0
K=N-J+1
M = K+1
DO 519 KK=M,N
519 T = T+C(K,KK)*Y(KK)
DO 520 KK=M,N
520 Y(KK) = Y(KK) - 2.*T*C(K,KK)
521 CONTINUE
T=0,0
DO 523 J = 1,N
523 T = T + Y(J)**2
XNORM=SQRT(T)
DO 524 J = 1,N
524 V(J,I) = Y(J) / XNORM
600 CONTINUE
700 RETURN
END

```

```

SUBROUTINE ORIENT(E,H,M,X,O,KOUNT,ENB,EV)
DIMENSION E(45,45),H(45,45),O(45,45),X(45),EV(45)
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
DO 170 IX=1,M
170 X(IX)=0
      WRITE (LUT,100)
      DO 200 I=1,M
      DO 200 J=1,M
200 X(I) = X(I) + H(I,J)*E(J,KOUNT)
      AQ=0,
      DO 230 I=1,M
230 AQ = AQ + X(I)*E(I,KOUNT)
      DO 260 I=1,M
      X(I)=0.
      DO 260 J=1,M
260 X(I) = X(I) + O(I,J)*E(J,KOUNT)
      AP=0,
      DO 271 I=1,M
271 AP = AP + E(I,KOUNT)*X(I)
      HLAMB = EV(KOUNT)
      CHISQ=ENB*(AP+HLAMB+AC/HLAMB-2.)
      CHISQ = ABS(CHISQ)
      K=M-1
      WRITE (LUT,280) KOUNT,CHISQ,K
      RETURN
C
100 FORMAT (33H0      VECTOR      CHISQUARE      DF   )
280 FORMAT (6X,I2, 9X,F9.3,4X,I3)
END

```

```

SUBROUTINE HETDST
COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
1 T(45),X(45),Z(45),NA,NB,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
2 MATPOS,KREYD
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
PROGRAM HETDST
ANDERSON AND BAHADURS GENERALIZED DISTANCE (R.REYMENT, JULY, 1966)
THE PROGRAM COMPUTES THE GENERALIZED DISTANCE FOR UNEQUAL
COVARIANCE MATRICES ACCORDING TO ANDERSON/BAHADUR(1962).

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INPUT = ON FIRST CARD,,,NC, VARIABLES, NO SPECIMENS IN SAMPLE A,  
AND NO. SPECIMENS IN SAMPLE B...INDIVIDUALS OF A AND INDIVIDUALS

OUTPUT = COMPUTED COVARIANCE MATRICES,SAMPLE SIZES,FOR EACH  
ITERATION,THE NUMBER OF ITERATION,VALUE OF ITERATION FACTOR T,  
COMPUTED Z DERIVING FROM THIS,INVERSE MATRIX AND ITS DETERMINANT,  
THE ESTIMATION VECTORS,INTERMEDIATE VALUES FOR DISTANCE,DISTANCE,  
SQUARE OF DISTANCE.

WRITE (LUT,1)

ENA=NA

ENB=NB

WRITE (LUT,18) NA,NB

REQUIRES COVARIANCE MATRICES,A,B,A VECTOR,SB(I) TO BE DETERMINED.

IT=0

K=1

T(K)=0.5

NV=M

HETDST

```

C      VECTOR SC CONTAINS THE DIFFERENCES OF MEANS
      WRITE (LUT,993)
      DO 997 IX=1,NV
997  WRITE (LUT,304) SC(IX)
      WRITE (LUT,733)
      DO 69 KK=1,4
      CALL BAHADU
      IF (KK.EQ.4) GO TO 720
      I=1,+(Z(K)*10.)
      IF (I.EQ. 1) GO TO 161
      K=K+1
      IF (I.GT. 1) GO TO (162,2001,2003),KK
      IF (KK-2) 30,31,32
30   T(K)=0.75
      GO TO 69
31   T(K)=0.675
      GO TO 69
32   T(K)=1.0
      GO TO 69
2003  T(K)=.8125
69    CONTINUE
2001  T(K)=.625
      DO 2004 KK=1,2
      CALL BAHADU
      IF (KK.EQ.2) GO TO 720
      I=1,+(Z(K)*10.)
      IF (I.EQ. 1) GO TO 161
      K=K+1
      IF (I.GT. 1) GO TO 2005
      T(K)=.6875
      GO TO 2004
2005  T(K)=.5625
2004  CONTINUE
162   T(K)=.25
      DO 2169 KK=1,3
      CALL BAHADU
      IF (KK.EQ.3) GO TO 720
      I=1,+(Z(K)*10.)
      IF (I.EQ. 1) GO TO 161
      K=K+1
      IF (I.GT. 1) GO TO (736,2007),KK
      T(K)=.375
      IF (KK.EQ.2) T(K)=.4375
      GO TO 2169
2007  T(K)=.3125
2169  CONTINUE
736   T(K)=.125
      DO 2008 KK=1,2
      CALL BAHADU
      IF (KK.EQ.2) GO TO 720
      I=1,+(Z(K)*10.)
      IF (I.EQ. 1) GO TO 161
      K=K+1
      IF (I.GT. 1) GO TO 2009
      T(K)=.1875
      GO TO 2008
2009  T(K)=.0625
2008  CONTINUE
720   IT = K-1
      I=1,+(Z(K)*10.)

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```

IF (I-1) 2710,161,2711
2710 SIGNTM=1.0
      GO TO 2703
2711 SIGNTM=-1.0
2703 IT=IT+1
      K=K+1
      DIFF(K)=ABS(T(K-1)-T(K-2))
      IF (DIFF(K) .LE. DIFF(K-1)) GO TO 2705
      T(K)=T(K-1)+SIGNTM*DIFF(K)/FLOAT(IT-2*IT/3)
2706 IF (DIFF(K) .LT. 0.00001) GO TO 161
      CALL BAHADU
      IF (IT-19) 720,720,169
2705 T(K)=T(K-1)+SIGNTM*DIFF(K)*0.5
      GO TO 2706
169 WRITE (LUT,184)
C   161 WRITE (LUT,174)
      DO 735 I=1,NV
735 WRITE (LUT,734) SB(I)
      IF(MATRIX,EQ.1) CALL EXIT
      CALL JZUP2
      CALL HAFCOV
      CALL REYDST
      IF(LDISCR.EQ.1) CALL DISTFN
C   REYDST COMPUTES T SQUARE AND VARIANCE RATIO FOR HET MATRICES
      RETURN
C
C
1 FORMAT(81H1 STATISTICAL DISTANCE FOR HETEROGENEOUS COVARIANCE MATR
1ICES )
      HETDSI
18 FORMAT(8H0  MA =,I3,7H  NB =,I3/3H0  ITERATIONS FOR DISTANCE CA
1LCULATION)
174 FORMAT(63H0ESTIMATE OF DISCRIMINANT VECTOR )
184 FORMAT(19H0WOULD NOT CONVERGE)
304 FORMAT(1X,7F12.5)
733 FORMAT(48H0ITERATION    Z          D          DSQ          DIFF T      )
734 FORMAT(2X,F10.5)
993 FORMAT(31H  SAMPLE DIFFERENCE MEAN VECTOR)
      END

SUBROUTINE BAHADU
DIMENSION ASC(45)
COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
1 T(45),X(45),Z(45),MA,NE,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
2 MATPOS,KREYD
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
C   SC IS DIFFERENCE OF MEANS
      NV=M
      TNEW=T(K)
      DO 151 I=1,NV
      DO 151 J=1,NV
151 C(I,J)=TNEW*A(I,J)+(1.-TNEW)*B(I,J)
      CALL MATINV (C,NV,ASC,0,DETERM,MW)
C   SOLUTION OF EQUATION FOR VECTOR SB(I) THRU MATINV AND MEAN VECTOR
      DO 153 I=1,NV
      SB(I)=0,
      DO 153 J=1,NV
153 SB(I)=C(I,J)*SC(J)*SB(I)
C   THIS GIVES THE VECTOR ESTIMATION DESIRED

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C      DISCRIMINANT VECTOR IS SB(I)
DO 154 I=1,NV
DO 154 J=1,NV
154 C(I,J)=A(I,J)*TNEW**2-((1.-TNEW)**2)*B(I,J)
DO 155 I=1,NV
X(I)=0,
DO 155 J=1,NV
155 X(I)=X(I)+C(I,J)*SB(J)
ZNEW=0,
DO 156 I=1,NV
156 ZNEW=ZNEW+X(I)*SB(I)
C      EXPANSION OF THE QUADRATIC FORM
Z(K)=ZNEW
TNEW=0,
Q=0,0
DO 167 I=1,NV
X(I)=0,0
167 Q=Q+SC(I)*SB(I)
Q=2,0*Q
DO 168 I=1,NV
ASC(I)=0,
DO 168 J=1,NV
168 ASC(J)=ASC(I)+A(I,J)*SB(J)
R=0,0
DO 170 I=1,NV
170 R=R+ASC(I)*SB(I)
R=SQRT(ABS(R))
DO 171 I=1,NV
ASC(I)=0,
DO 171 J=1,NV
171 ASC(I)=ASC(I)+B(I,J)*SB(J)
AS=0,0
DO 172 I=1,NV
172 AS=AS+ASC(I)*SB(I)
AS=SQRT(ABS(AS))
U=Q/(R+AS)
DSQ=D**2
WRITE (LUT,159) IT,Z(K),D,DSQ,DIFF(K)
RETURN

C
159 FORMAT(2X,I2,7X,F8.5,2X,F6.3,3X,F6.3,3X,F10.7)
END

SUBROUTINE HAFCOV
COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
1 T(45),X(45),Z(45),M,NA,NE,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
2 MATPOS,KREYD
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
WRITE (LUT,407)
CALL GETMAT(A,1)
CALL GETMAT(B,2)
DO 408 I=1,M
DO 408 J=1,M
408 C(I,J)=(A(I,J)+B(I,J))/2,
WRITE (LUT,200)
CALL MATOUT(C)
CALL MATINV(C,M,Z,0,DETERM,MW)
DO 410 I=1,M

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Z(I)=0.
DO 410 J=1,M
410 Z(I) = Z(I) + C(I,J)*SC(J)
DSQ=0.
DO 411 I=1,M
411 DSQ = DSQ + SC(I)*Z(I)
DSQ=ABS(DSQ)
DS=SQRT(DSQ)
WRITE (LUT,222) DS,DSQ
RETURN

C
C
200 FORMAT (24H0 ROW COVARIANCE MATRIX )
222 FORMAT( 33H0 DISTANCE, DISTANCE SQUARED /7X,F7.3,8X,F9.3/)
407 FORMAT(61H1 HETEROGENEUS DISTANCE BY AVERAGING COVARIANCE MATR
1ICES )
END

SUBROUTINE REYDST
DIMENSION XA(45),XB(45),SLM(45),DISC(45),XC(45)
COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
1 T(45),X(45),Z(45),M,NA,NE,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
2 MATPOS,KREYD
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
C----- RETURN IF KREYD IS LE 0
IF (KREYD ,LE, 0) RETURN
WRITE (LUT,4165)
DSUM = 0.
ENB = NB
ENA = NA
NAB = NA
TK = M
AN = SQRT(ENB/ENA)
WRITE (LUT,3) AN
C      REMEMBER THAT A IS THE LARGER SAMPLE AND NA IS GRT TH NB
C      REMEMBER THAT VECTOR SC QCONTAINS DIFFERENCE OF MEANS
C***** ****
C----- THE STATEMENTS FRUM DC 6008 UNTIL 200 CONTINUE ARE MACHINE
C----- DEPENDENT. THEY USE DC 3600-FORTRAN STATEMENTS TO STORE SAMPLE
C----- A AND SAMPLE B ON A RANDOM ACCESS DRUM STORAGE UNIT
C----- IF YOU HAVE NO RANDOM ACCESS UNIT YOU CAN SKIP THIS STATEMENTS
C----- AND MAKE SOME CHANGES IN ROUTINE RANKRT.
REWIND MTARA
REWIND MTA
DO 6008 IV = 1,NAB
READ (MTA) (XA(IX),IX=1,M)
BUFFER OUT (MTARA,1) (XA(1),XA(M))
6011 IF (UNIT,MTARA) 6011,6008
6008 CONTINUE
C----- SAMPLE A STORED NOW ON DRUM IN RANDOM ACCESS
REWIND MTBRA
REWIND MTB
DO 200 IV=1,NE
READ (MTB) (XB(IX),IX=1,M)
BUFFER OUT (MTBRA,1) (XB(1),XB(M))
205 IF (UNIT,MTBRA) 205,200
200 CONTINUE
C----- SAMPLE B IS NOW STORED ON DRUM IN RANDOM ACCESS
C***** ****
6009 NAB = NB

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```

DO 500 KK=1,KREYD
DO 6050 I=1,M
XC(I) = 0.
6050 X(I) = U.
DO 6007 I=1,M
SUM(I) = 0.
DO 6007 J=1,M
6007 A(I,J) = 0.
CALL CLEAR
DO 40 IV=1,NAB
CALL RANKRT(XA,0)
CALL RANKRT(XE,1)
DO 20 I=1,M
20 XC(I) = XA(I)*AN
DO 22 I=1,M
22 X(I) = ABS(XB(I)-XC(I))
C NEW OBSERVATIONS STORED IN X(I)
DO 23 I=1,M
23 SUM(I) = SUM(I) + X(I)
DO 25 I=1,M
DO 25 J=1,M
25 A(I,J) = A(I,J) + X(I)*X(J)
40 CONTINUE
DO 45 I=1,M
DO 45 J=1,M
45 A(I,J)=A(I,J) - SUM(I)*SUM(J)/ENB/(ENB-1.0)
COVARIANCE MATRIX OF THE NEW VARIABLES
CALL MATINV(A,M,B,0,DETERM,MW)
DO 60 I=1,M
DISC(I)=0.
DO 60 J=1,M
60 DISC(I) = DISC(I) + A(I,J)*SC(J)
DS = 0.
DO 70 I=1,M
70 DS = DS + DISC(I)*SC(I)
DS = DS*2.0
DSUM = DSUM + DS
TEST = DSUM/FLGAT(KK)
WRITE (LUT,76) KK,DS
500 CONTINUE
WRITE (LUT,80)
DS=DSUM/FLOAT(KREYD+1)
TT = DS*ENB/2.
WRITE (LUT,85) DS,TT
F = (TT*(ENB-FLOAT(M)))/((ENB-1.)*FLOAT(M))
NAA = NB-1
WRITE (LUT,95) F,NAA,M
RETURN
C
3 FORMAT(24H0 MULTIPLICATION FACTOR /8H0 AN = ,F14.7/32H0 ITERATI
1ON REMENTS DSQ TEST )
76 FORMAT(6X,I2,8X,F14.4,8X)F14.4)
80 FORMAT(28H0 SIGNIFICANCE COMPUTATIONS /1H0)
85 FORMAT(25H0 AVE,DSQ TSG /2(3X,F10.3))
95 FORMAT(23H0 VARIANCE RATIO IS = F14.2,4H FOR 14,5H AND 12,19H DEGR
1EES OF FREEDOM )
4165 FORMAT(30H1 HETEROGENEUS SIGNIFICANCES      )
END

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SUBROUTINE CLEAR
C----- CLEAR IS A ROUTINE WHICH MUST BE CALLED BEFORE FIRST CALL RANKRT
C----- IT CLEARS THE INDICATOR ARRAY KAT.
COMMON /RANK0/ KAT(500)
DO 40 I=1,500
40 KAT(I)=0
RETURN
C***** IF YOU HAVE NO RANDOM ACCESS UNIT YOU MUST ADD THE FOLLOWING
C----- CODE,
C COMMON /UNITS/ ...
C REWIND MTA
C REWIND MTB
C***** END

SUBROUTINE RANKRT(XD,KAESK)
C----- RANKRT PICKS RANDOMLY ONE RECORD FROM SAMPLE A (IF KABSW=0 )
C----- OR FROM SAMPLE B ( IF KABSW=1 ), THIS VERSION ASSUMES THAT
C----- SAMPLE A AND SAMPLE B IS PLACED EARLIER ON RANDOM ACCESS UNITS
C----- MTARA RESP MTBRA. BEFORE FIRST CALL RANKRT THE INDICATOR ARRAY
C----- KAT MUST BE CLEARED BY CALL CLEAR.
C----- VERSION 2, WHICH CAN HANDLE TWO PARALLEL RANDOM FILES
DIMENSION XD(45)
COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
1 T(45),X(45),Z(45),M,NA,NE,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
2 MATPOS,KREYD
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
COMMON /RANK0/ KAT(500)
KAB=KABSW
IF (KAB) 45,45,47
C----- PICK UP A CARD FROM UNIT MTARA
45 NT=NA
LUN=MTARA
GO TO 48
C----- PICK UP A CARD FROM UNIT MTBRA
47 NT=NB
LUN=MTBRA
48 CONTINUE
C----- GET A RANDOM NUMBER FROM 1 TO NT+1
C----- X=RANF(-1) WILL GIVE A NEW FLOATING POINT RANDOM NUMBER IN THE
C----- INTERVAL 0.0 UNTIL 1.0 EVERY TIME RANF IS CALLED.
20 I=RANF(-1)*FLOAT(NT)+1.0
C----- DON'T ACCEPT I .GT. NT
IF (I .GT. NT) GO TO 20
C----- GET INDICATOR FOR UNIT MTBRA
IK=KAT(I)/10
C----- GET INDICATOR FOR UNIT MTARA
IR=KAT(I)-IK*10
C----- TEST IF WE JUST NOW ARE USING MTARA
IF (KAB .EQ. 0) IK=IK
C----- IF THIS CARD ALREADY USED TRY TO GET A NEW RANDOM NUMBER
IF (IK ,NE, 0) GO TO 20
C----- POSITION THE DRUM
CALL LOCATE (LUN,(I-1)*N)
BUFFER IN (LUN,1) (XU(1),XD(M))
11 IF (UNIT,LUN) 11,12
C----- INDICATE THAT THIS CARD IS USED
12 KAT(I)=KAT(I)+KAB*10
RETURN

```

```

C***** IF YOU HAVE NO RANDOM ACCES UNIT YOU CAN USE THE FOLLOWING CODE
C----- COMMON , , , ,
C----- DIMENSION , , , ,
C----- DIMENSION LUNPOS(2)
C----- DATA (LUNPOS=0,0)      OR      DATA LUNPOS/0,0/
C----- KAB=KABSW
C----- IF (KAB) 45,45,47
C 45 NT=NA
C     LUN#MTA
C     GO TO 48
C 47 NT=NB
C     LUN#MTB
C 48 CONTINUE
C 20 I=RANF(-1)*FLOAT(NT)+1.0
C     IF (I .GT. NT) GO TO 20
C     IK=KAT(I)/10
C     IR=KAT(I)-IK*10
C     IF (KAB.EQ. 0) IK=IR
C     IF (IK .NE. 0) GO TO 20
C     J=I-1-LUNPOS(KAB+1)
C     IF (J) 20,60,70
C 50 J=-J
C     DO 51 J1=1,J
C 51 BACKSPACE LUN
C     GO TO 60
C 70 DO 71 J1=1,J
C 71 READ(LUN) J2
C 60 LUNPOS(KAB+1)=I
C     READ(LUN) (XD(I)X),IX=1,M)
C 12 KAT(I)=KAT(I)+KAB*10
C     RETURN
C***** END

```

```

SUBROUTINE DISTFN
COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
1 T(45),X(45),Z(45),M,NA,NE,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
2 MATPOS,KREYD
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
ALL INPUT OF DATA IS DONE IN CALLING PROGRAM
WRITE (LUT,300)
ENA=NA
ENB=NB
NDF=NA+NB-2
WRITE (LUT,200) NA,NB,NDF
DO 13 I=1,M
 13 WRITE (LUT,272) I,SA(I),SE(I),SC(I)
4448 CONTINUE
CALL GETMAT (C,3)
CALL MATINV(C,M,Z,0,DETERM,MW)
DO 17 I=1,M
  Z(I)=0.
  DO 17 J=1,M
    17 Z(I) = Z(I) + C(I,J)*SC(J)
    WRITE (LUT,309)
    WRITE (LUT,222) (Z(I),I=1,M)
    DO 1064 I=1,M
1064 SB(I) = Z(I)

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```

DSQ=0.
DO 18 I=1,M
18 DSQ = DSQ + SC(I)*Z(I)
DSQ=ABS(DSQ)
WRITE (LUT,310)
EM=M
DS=SQRT(DSQ)
WRITE (LUT,223) DS,DSQ
WRITE (LUT,312)
DSQ=(ENA*ENB/(ENA+ENB))*DSQ
WRITE (LUT,313) DSQ
DSQ=DSQ*((ENA+ENB-EM-1.)/(EM*(ENA+ENB-2.)))
WRITE (LUT,314) DSQ
NP = NA+NB-M-1
WRITE (LUT,315) M,NP
CALL JZUP2
RETURN
C
C
200 FORMAT(13H      N1 8,13,12H      N2 #13,12H      DF =,I3/
151H VAR      MEAN VECTOR 1  MEAN VECTOR 2  DIFF,VECTOR )
222 FORMAT(2X,12F10.5)
223 FORMAT(2X,F10.5,11X,F10.5)
272 FORMAT(14,F14.3,2F15.3)
300 FORMAT(32H1      STATISTICAL DISTANCE ,//1H0)
309 FORMAT(28H0 DISCRIMINATOR COEFFICIENTS)
310 FORMAT(63H0 RESULTS FOR MAHANALOBIS DSQUARE /33H      D
1      D SQUARE )
312 FORMAT(40H0 SIGNIFICANCE FOR D SQUARE AND T SQUARE/)
313 FORMAT(11H0T SQUARE #F13.3)
314 FORMAT( 4H0F #F13.2)
315 FORMAT(6H0DF1 = I3,7F  DF2 = I3  )
END

SUBROUTINE MATINV(A,N,B,M,DETERM,JW)
DIMENSION IPIVOT(45),A(JW,JW),B(JW,1),INDEX(45,2),PIVOT(45)
EQUIVALENCE(IROW,JROW),(ICOLUMN,JCOLUMN),(AMAX,T,SWAP)
DETERM=1.
DO 20 J=1,N
20 IPIVOT(J)=0
DO 550 I=1,N
AMAX=0.
DO 105 J=1,N
IF(IPIVOT(J)-1)60,105,60
60 DO 100 K=1,N
IF(IPIVOT(K)-1)80,100,740
80 IF(ABS(AMAX)-ABS(A(J,K)))85,100,100
85 IROW=J
ICOLUMN=K
AMAX=A(J,K)
100 CONTINUE
105 CONTINUE
IPIVOT(ICOLUMN)=IPIVOT(ICOLUMN)+1
C     INTERCHANGE ROWS TO PUT PIVOT ELEMNT ON DIAGONAL
IF(IROW-ICOLUMN)140,260,140
140 DETERM=DETERM
DO 200 L=1,N
SWAP=A(IROW,L)
A(IROW,L)=A(ICOLUMN,L)
200 A(ICOLUMN,L)=SWAP

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```

IF(M)260,260,210
210 DO 250 L=1,N
SWAP=B(IROW,L)
B(IROW,L)=B(ICOLUMN,L)
250 B(ICOLUMN,L)=SWAP
260 INDEX(1,1)=IROW
INDEX(1,2)=ICOLUMN
PIVOT(I)=A(ICOLUMN,ICOLUMN)
DETERM=DETERM*PIVOT(I)
C DIVIDE PIVOT ROW BY PIVOT ELEMENT
A(ICOLUMN,ICOLUMN)=1,
DO 350 L=1,N
350 A(ICOLUMN,L)=A(ICOLUMN,L)/PIVOT(I)
IF(M)380,380,360
360 DO 370 L=1,M
370 B(ICOLUMN,L)=B(ICOLUMN,L)/PIVOT(I)
380 DO 550 L1=1,N
IF(L1-ICOLUMN)400,550,400
400 T=A(L1,ICOLUMN)
A(L1,ICOLUMN)=0,
DO 450 L=1,N
450 A(L1,L)=A(L1,L)-A(ICOLUMN,L)*T
IF(M)550,550,460
460 DO 500 L=1,M
500 B(L1,L)=B(L1,L)-B(ICOLUMN,L)*T
550 CONTINUE
DO 710 I=1,N
L=N+1-I
IF(INDEX(L,1)-INDEX(L,2))630,710,630
630 JROW=INDEX(L,1)
JCOLUMN=INDEX(L,2)
DO 705 K=1,N
SWAP=A(K,JROW)
A(K,JROW)=A(K,JCOLUMN)
A(K,JCOLUMN)=SWAP
705 CONTINUE
710 CONTINUE
740 RETURN
END

```

SUBROUTINE PLTHS1(XX,III,START,ENDX,NUM)

PLOTS A FREQUENCY DISTRIBUTION IN NUM INCREMENTS FOR THE VALUES IN PLTH  
 THE ARRAY XX BETWEEN START AND ENDX IN VSUVE. THE FIRST III PLTH  
 VALUES IN THE XXX ARHAY ARE CONSIDERED. PLTH

DIMENSION XX(100),KCNT(100),LINE(101) PLTHS102

DATA (LB=1H),(LAST=1H\*),(LDASH=1H=),(LI=1HI)

C=F IV- DATA LB,LAST,LDASH,LI /1H ,1H\*,1H-,1H1/

COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA

DO 1 I=1,NUM

1 KONT(I) = 0

C SET NUM COUNTERS TO ZERO,

XNUM = NUM

XINT = (ENDX-START)/XNUM

C THE INTERVAL SIZE FOR THE NUM INTERVALS IS XINT.

NUM2 = 100/NUM

NUM3 = NUM2 - 1

PLTHS104

PLTHS105

PLTH

PLTHS106

PLTHS107

PLTH

PLTHS108

PLTHS109

```

NUM4 = NUM2*NUM+1
DO 5 I=1,101
5 LINE(I) = LB
C LINE IS SET EQUAL TO BLANKS.
LINE(NUM4) = LI
C THE LAST PLACE TO BE USED IN LINE IS SET EQUAL TO AN LI
DO 10 I=1,III
C LOOK AT III POINTS IN THE XX ARRAY.
IF(XX(I).LT.START.OR.XX(I).GT.ENDX) GO TO 10
C IF XX(I) IS OUTSIDE THE RANGE IGNORE IT.
KNT = (XX(I)-START)/XINT
KNT = KNT + 1
C KNT IS THE INTERVAL THAT XX(I) FALLS INTO?
IF(KNT.GT.NUM) KNT = NUM
C IF XX(I) = ENDX KNT COULD BE GREATER THAN NUM.
KONT(KNT) = KONT(KNT) + 1
C INCREASE THE KNT COUNTER BY 1.
10 CONTINUE
C
KMAX = KONT(1)
DO 20 I=2,NUM
IF(KONT(I).GT.KMAX) KMAX = KONT(I)
C FIND THE VALUE OF MAXIMUM COUNTER.
20 CONTINUE
C
KSCALE = KMAX/50
C KSCALE IS THE VERTICAL SCALE FACTOR.
IF(KSCALE*50.NE.KMAX) KSCALE = KSCALE + 1
WRITE (LUT,100) XINT,START,ENDX,NUM
NTOP = KMAX + KSCALE
DO 21 I=1,NUM
IF(KONT(I).EQ.0) KONT(I) = -KSCALE
21 CONTINUE
30 NTOP = NTOP-KSCALE
IF(NTOP.LE.(-KSCALE)) GO TO 40
IF(NTOP.LT.0) NTOP = 0
ICONT = 0
DO 32 I=1,NUM
IF(KONT(I).LT.NTOP) GO TO 32
C IF THE COUNTER FOR THE INTERVAL IS GREATER THAN OR EQUAL TO NTOP
C PUT ASTERISKS INTO THE PART OF LINE ALLOCATED TO THIS INTERVAL.
DO 33 I=1,NUM2
C THERE ARE NUM2 SPACES IN LINE FOR EACH INTERVAL.
INX = I + ICONT
33 LINE(INX) = LAST
32 ICONT = ICONT + NUM2
IF(NTOP.EQ.0) GO TO 40
WRITE (LUT,101) NTOP,LINE
GO TO 30
C
40 IF(NUM3.EQ.0) GO TO 50
C THIS SEGMENT FROM HERE TO RETURN PRINTS THE BASE LINE WITH I AT
C THE END OF EACH INTERVAL AND MINUSES BETWEEN.
ICONT = 0
DO 41 I=1,NUM
DO 42 I=1,NUM3
INX = I + ICONT
42 LINE(INX) = LDASH
LINE(INX+1) = LI
41 ICONT = ICONT + NUM2

```

PLTHSI11  
PLTHSI12  
PLTH  
PLTH  
PLTHSI14  
PLTH  
PLTHSI15  
PLTHSI16  
PLTHSI17  
PLTH  
PLTHSI18  
PLTH  
PLTHSI19  
PLTH  
PLTHSI20  
PLTH  
PLTHSI21  
PLTHSI22  
PLTHSI23  
PLTH  
PLTHSI24  
PLTHSI25  
PLTH  
PLTHSI26  
PLTHSI27  
PLTHSI28  
PLTHSI29  
PLTHSI30  
PLTHSI31  
PLTHSI32  
PLTHSI33  
PLTHSI34  
PLTHSI35  
PLTHSI36  
PLTHSI37  
PLTH  
PLTH  
PLTHSI38  
PLTH  
PLTHSI39  
PLTHSI40  
PLTHSI41  
PLTHSI43  
PLTHSI44  
PLTH  
PLTH  
PLTHSI45  
PLTHSI46  
PLTHSI47  
PLTHSI48  
PLTHSI49  
PLTHSI50  
PLTHSI51

```

      WRITE (LUT,102) LINE
      RETURN
      PLTHSI53
      PLTH
      PLTHSI54
      PLTHSI55
      PLTHSI57

C
      50 DO 51 I=1,NUM
      51 LINE(I) = LI
      WRITE (LUT,102) LINE
      RETURN
      PLTHSI57

C
      100 FORMAT(34H1      HISTOGRAM FOR DISCRIMINANTS /14HO INTERVAL # ,
      1F10.4,16H DATA STARTS AT ,F10.4/11HO ENDS AT ,F10.4,19H CLASS INT
      2ERVALS = ,I4)
      101 FORMAT(15,2X,1HI,101A1)
      102 FORMAT(7X,1HI,101A1)
      END
      PLTHSI61
      PLTHSI62

      SUBROUTINE DISTRB
C THIS SUBROUTINE TESTS THE INDIVIDUAL VARIABLES FOR DEVIATIONS
C FROM THE UNIVARIATE NORMAL DISTRIBUTION BY CALCULATING THE
C COEFFICIENTS OF SKEWNESS AND KURTOSIS AND THEIR STANDARD ERRORS
      DIMENSION STDEVA(45),STDEVB(45),R(45,4)
      COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
      1 T(45),X(45),Z(45),NA,NE,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
      2 MATPOS,KREYD
      COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMAEC,MTARA,MTBRA
      DO 5 I=1,M
      STDEVA(I)=SQRT(A(I,I))
      5 STDEVB(I) = SQRT(B(I,I))
      WRITE (LUT,6)
      INDKS = 1
      WRITE (LUT,10) INDKS
      DO 14 J=1,M
      14 WRITE (LUT,15) SA(J),STDEVA(J)
      INDKS = 2
      WRITE (LUT,10) INDKS
      DO 19 J=1,M
      19 WRITE (LUT,15) SB(J),STDEVB(J)
      KTOT = NA
      KNT=0
C----- SAMPLE A FIRST
      LUN=MTA
      DO 24 I=1,M
      DO 24 J=1,4
      24 R(I,J)=0.
      251 REWIND LUN
      25 CONTINUE
      READ (LUN) (X(I),I=1,M)
      TOT=KTOT
      ERROR1=SQRT((6.*TOT*(TOT+1.))/((TOT-2.)*(TOT+1.)*(TOT+3.)))
      ERROR2=SQRT((24.*TOT*(TOT+1.)**2)/((TOT-3.)*(TOT-2.)*(TOT+3.)*(TOT
      1+5.)))
      KNT = KNT+1
      DO 30 I=1,M
      PRODX=1.
      DO 30 J=1,4
      PRODX=PRODX*X(I)
      30 R(I,J) = R(I,J) + PRODX
      IF(KNT-KTOT) 25,35,39
      35 DO 40 J=1,4
      DO 40 I=1,M

```

```

40 R(I,J) = R(I,J)/FLOAT(KTOT)
  WRITE (LUT,45)
  DO 100 I=1,M
    XM3=R(I,3)-3.*R(I,1)*R(I,2)+2.*R(I,1)**3
    XG1=XM3/(STDEVA(I)**3)
    XM4=R(I,4)-4.*R(I,1)*R(I,3)+6.*R(I,1)**2*R(I,2)-3.*R(I,1)**4
    XG2=XM4/(STDEVA(I)**4)=3,
    XXG1 = XG1
    XXG2 = XG2
    XXG1 = XXG1/ERROR1
    XXG2 = XXG2/ERROR2
    WRITE (LUT,46) I,XG1,XXG1,XG2,XXG2
100 CONTINUE
  WRITE (LUT,52) ERRQR1,ERRCR2
  DO 55 I=1,M
  DO 55 J=1,4
55 R(I,J)=0.
  DO 54 I=1,M
54 STDEVA(I) = STDEVB(I)
  IF(KTOT.EQ.NA) GO TO 56
  IF(KTOT.EQ.NB) GO TO 200
56 WRITE (LUT,60)
  KNT = 0
C----- SAMPLE B THE SECOND TIME
  KTOT=NB
  LUN=MTB
  GO TO 251
200 WRITE (LUT,6)
  RETURN

```

```

C
C
6 FORMAT(44H1  DISTRIBUTION DETAILS FOR REFERENCE SAMPLE      )
10 FORMAT(33H0  MEANS FOR SAMPLE I1,46H                      ) ST
1  STANDARD DEVIATIONS FOR SAMPLE I1)
15 FORMAT(15X,F18.5,12X,F18.5)
45 FORMAT(52H0  VARIABLE SKEWNESS   T(SKEW) KURTOSIS   T(KURT)  )
46 FORMAT(6X,I2,4X,4F10.4)
52 FORMAT(25H0  ST.DEV,_CR SKEWNESS = ,F14.7/25H0  ST.DEV+FOR KURTOSI
15 = ,F14.7)
60 FORMAT(45H1  DISTRIBUTION DETAILS FOR COMPARISON SAMPLE)
END

```

```

SUBROUTINE JZUP2
DIMENSION ODATA(1000),BIG(3),SMALL(3)
COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
1 T(45),X(45),Z(45),M,NA,NE,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
2 MATPOS,KREYD
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
NTOTAL=NA+N
DO 5 I=1,NTOTAL
  5 ODATA(I) = 0.
  DO 45 IREP = 1,2
    IF (IREP .EQ. 2) GO TO 30
C----- SAMPLE A
  LUN=MTA
  JA=1
  NE=NA
  GO TO 31
C----- SAMPLE B

```

```

30 LUN=MTB
   JA=NA+1
   NEJ=NTOTAL
31 REWIND LUN
   DO 10 II=JA,NEJ
      READ (LUN) (X(I),I=1,M)
      DO 9 IV=1,M
         9 ODATA(II) = SB(IV)*X(IV) + ODATA(II)
10 CONTINUE
45 CONTINUE
   NLIM = NTOTAL
   LOWER=1
   DO 29 IX=1,3
      GO TO (62,60,61),IX
60 IF (INVST(2),EQ,INVST(1)) GO TO 29
   GO TO 62
61 IF (INVST(3),EQ,INVST(1).OR.INVST(3) .NEQ. INVST(2)) GO TO 29
62 BIG(IX)=ODATA(LOWER)
   SMALL(IX)=ODATA(LOWER)
   DO 22 IV = LOWER,NLIM
      IF(ODATA(IV).GT.BIG(IX)) BIG(IX)=ODATA(IV)
      IF(ODATA(IV).LT.SMALL(IX)) SMALL(IX)=ODATA(IV)
22 CONTINUE
   GO TO 20
20 CALL PLTHS1(ODATA(LOWER),NTOTAL,SMALL(IX),BIG(IX),INVST(IX))
29 CONTINUE
55 RETURN
END

```

SPCL17/22  
SPCL17/24

SPCL17/30

SPCL17/34  
SPCL17/36

SPCL17/38  
SPCL17/40

SPCL17/60

```

SUBROUTINE MATOUT (STMAT)
C----- MATOUT PRINTS OUT THE M*M MATRIX STMAT ON THE PRINTER UNIT LUT
DIMENSION STMAT(45,45)
COMMON A(45,45),B(45,45),C(45,45),SA(45),SB(45),SC(45),DIFF(45),
1 T(45),X(45),Z(45),M,NA,NE,K,MW,MATRIX,LDISCR,LOGDEC,INVST(3),IT,
2 MATPOS,KREYD
COMMON /UNITS/ LIN,LUT,MTA,MTB,MTMABC,MTARA,MTBRA
C----- MATOUT PRINTS OUT THE MATRIX STMAT ON LOGICAL UNIT LUT.
MEND=0
1 MBEG=MEND+1
MEND=MBEG+9
C----- PRINT MAXIMUM 10 COLUMNS EACH TIME
IF (MEND ,GT, M) MEND=M
WRITE (LUT,101) (J,J=MBEG,MEND)
DO 2 I=1,M
2 WRITE (LUT,102) I,(STMAT(I,J),J=MBEG,MEND)
C----- IF ALL M COLUMNS HAVE BEEN PRINTED THEN RETURN
IF (MEND ,EQ, M) RETURN
C----- OR ELSE PRINT NEXT 10 COLUMNS
GO TO 1
101 FORMAT (1H0,10I12)
102 FORMAT (2X,I2,2X,10F12.5)
END

```

APPENDIX B. -Example of Output from Program.

ANALYSIS OF HOMOGENEITY OF COVARIANCE MATRICES  
TEST OF NEW ORNTDIST FOR 45\*45 MATRICES, GOTLAND.

DATA CARDS READ WITH VARIABLE FORMAT (F9.2,3F10,2)

REFERENCE SAMPLE SIZE BASED ON POPULATION OF 128

COMPARISON SAMPLE SIZE IS 95

FIRST DATA INPUT SET

735.0000000	720.0000000	615.0000000	435.0000000
715.0000000	660.0000000	580.0000000	430.0000000
710.0000000	700.0000000	625.0000000	360.0000000
770.0000000	755.0000000	655.0000000	405.0000000
770.0000000	730.0000000	640.0000000	450.0000000
720.0000000	660.0000000	565.0000000	395.0000000
765.0000000	715.0000000	650.0000000	460.0000000
760.0000000	735.0000000	650.0000000	455.0000000
740.0000000	725.0000000	655.0000000	440.0000000
760.0000000	740.0000000	675.0000000	410.0000000
775.0000000	730.0000000	610.0000000	360.0000000
710.0000000	700.0000000	610.0000000	340.0000000
705.0000000	675.0000000	600.0000000	435.0000000
790.0000000	740.0000000	660.0000000	420.0000000
740.0000000	710.0000000	575.0000000	460.0000000
715.0000000	670.0000000	590.0000000	385.0000000
765.0000000	705.0000000	635.0000000	470.0000000
760.0000000	740.0000000	610.0000000	415.0000000
745.0000000	725.0000000	655.0000000	425.0000000
730.0000000	700.0000000	635.0000000	410.0000000
805.0000000	745.0000000	655.0000000	465.0000000
780.0000000	725.0000000	630.0000000	420.0000000
775.0000000	740.0000000	650.0000000	735.0000000
760.0000000	720.0000000	650.0000000	410.0000000
755.0000000	705.0000000	610.0000000	385.0000000
805.0000000	720.0000000	650.0000000	460.0000000
740.0000000	690.0000000	570.0000000	400.0000000
730.0000000	695.0000000	620.0000000	405.0000000
770.0000000	735.0000000	645.0000000	400.0000000
755.0000000	755.0000000	600.0000000	450.0000000
765.0000000	770.0000000	625.0000000	410.0000000
755.0000000	710.0000000	630.0000000	400.0000000
760.0000000	720.0000000	645.0000000	435.0000000
780.0000000	750.0000000	620.0000000	410.0000000
780.0000000	735.0000000	675.0000000	370.0000000
800.0000000	730.0000000	595.0000000	400.0000000
750.0000000	685.0000000	585.0000000	425.0000000
750.0000000	740.0000000	610.0000000	400.0000000
750.0000000	730.0000000	620.0000000	395.0000000
720.0000000	690.0000000	550.0000000	395.0000000
735.0000000	750.0000000	640.0000000	415.0000000
735.0000000	685.0000000	560.0000000	380.0000000
760.0000000	680.0000000	605.0000000	440.0000000
800.0000000	720.0000000	685.0000000	420.0000000
795.0000000	750.0000000	650.0000000	435.0000000
730.0000000	715.0000000	630.0000000	410.0000000

775,0000000	775,0000000	630,0000000	450,0000000
790,0000000	740,0000000	615,0000000	425,0000000
780,0000000	775,0000000	655,0000000	385,0000000
765,0000000	735,0000000	645,0000000	420,0000000
780,0000000	810,0000000	660,0000000	410,0000000
775,0000000	765,0000000	630,0000000	360,0000000
790,0000000	760,0000000	640,0000000	470,0000000
825,0000000	800,0000000	690,0000000	430,0000000
740,0000000	720,0000000	645,0000000	430,0000000
775,0000000	795,0000000	670,0000000	380,0000000
760,0000000	785,0000000	650,0000000	380,0000000
745,0000000	735,0000000	650,0000000	400,0000000
730,0000000	705,0000000	600,0000000	475,0000000
800,0000000	750,0000000	635,0000000	470,0000000
750,0000000	770,0000000	620,0000000	410,0000000
780,0000000	770,0000000	630,0000000	415,0000000
745,0000000	675,0000000	595,0000000	365,0000000
735,0000000	755,0000000	625,0000000	390,0000000
780,0000000	735,0000000	650,0000000	425,0000000
740,0000000	765,0000000	595,0000000	375,0000000
760,0000000	705,0000000	650,0000000	440,0000000
730,0000000	660,0000000	565,0000000	405,0000000
735,0000000	670,0000000	650,0000000	410,0000000
715,0000000	710,0000000	650,0000000	410,0000000
720,0000000	690,0000000	585,0000000	405,0000000
710,0000000	645,0000000	580,0000000	410,0000000
750,0000000	750,0000000	600,0000000	415,0000000
730,0000000	665,0000000	630,0000000	390,0000000
780,0000000	735,0000000	635,0000000	395,0000000
770,0000000	765,0000000	660,0000000	385,0000000
755,0000000	740,0000000	615,0000000	450,0000000
725,0000000	730,0000000	610,0000000	415,0000000
695,0000000	670,0000000	630,0000000	375,0000000
755,0000000	670,0000000	595,0000000	410,0000000
740,0000000	735,0000000	635,0000000	425,0000000
705,0000000	665,0000000	575,0000000	425,0000000
660,0000000	700,0000000	630,0000000	405,0000000
690,0000000	690,0000000	590,0000000	370,0000000
740,0000000	715,0000000	670,0000000	390,0000000
740,0000000	670,0000000	635,0000000	385,0000000
715,0000000	685,0000000	580,0000000	420,0000000
745,0000000	690,0000000	580,0000000	400,0000000
745,0000000	720,0000000	650,0000000	430,0000000
735,0000000	715,0000000	630,0000000	450,0000000
700,0000000	725,0000000	610,0000000	405,0000000
735,0000000	715,0000000	600,0000000	470,0000000
775,0000000	775,0000000	695,0000000	450,0000000
750,0000000	730,0000000	605,0000000	385,0000000
735,0000000	730,0000000	605,0000000	390,0000000
765,0000000	700,0000000	660,0000000	420,0000000
725,0000000	730,0000000	650,0000000	470,0000000
795,0000000	790,0000000	660,0000000	490,0000000
770,0000000	725,0000000	695,0000000	425,0000000
810,0000000	780,0000000	645,0000000	450,0000000
745,0000000	775,0000000	640,0000000	470,0000000
780,0000000	720,0000000	610,0000000	425,0000000
770,0000000	745,0000000	635,0000000	450,0000000
740,0000000	720,0000000	670,0000000	420,0000000
770,0000000	770,0000000	625,0000000	430,0000000
750,0000000	720,0000000	660,0000000	435,0000000

760.0000000	745.0000000	645.0000000	380.0000000
745.0000000	725.0000000	625.0000000	370.0000000
805.0000000	820.0000000	690.0000000	440.0000000
790.0000000	780.0000000	640.0000000	395.0000000
705.0000000	680.0000000	600.0000000	390.0000000
725.0000000	745.0000000	680.0000000	435.0000000
730.0000000	690.0000000	600.0000000	390.0000000
765.0000000	760.0000000	640.0000000	415.0000000
690.0000000	720.0000000	595.0000000	395.0000000
740.0000000	695.0000000	630.0000000	400.0000000
705.0000000	695.0000000	600.0000000	435.0000000
750.0000000	740.0000000	605.0000000	420.0000000
670.0000000	635.0000000	575.0000000	400.0000000
690.0000000	645.0000000	600.0000000	425.0000000
730.0000000	710.0000000	600.0000000	450.0000000
750.0000000	700.0000000	610.0000000	370.0000000
680.0000000	645.0000000	520.0000000	365.0000000
765.0000000	750.0000000	640.0000000	450.0000000
695.0000000	690.0000000	610.0000000	405.0000000
715.0000000	720.0000000	600.0000000	440.0000000
735.0000000	735.0000000	655.0000000	400.0000000
725.0000000	725.0000000	630.0000000	420.0000000

#### SECOND DATA INPUT SET

690.0000000	880.0000000	720.0000000	560.0000000
810.0000000	1025.0000000	870.0000000	535.0000000
720.0000000	890.0000000	820.0000000	580.0000000
830.0000000	1100.0000000	875.0000000	520.0000000
750.0000000	945.0000000	865.0000000	520.0000000
775.0000000	995.0000000	890.0000000	520.0000000
745.0000000	880.0000000	810.0000000	485.0000000
810.0000000	980.0000000	820.0000000	580.0000000
755.0000000	975.0000000	790.0000000	590.0000000
770.0000000	995.0000000	850.0000000	510.0000000
690.0000000	945.0000000	740.0000000	490.0000000
725.0000000	895.0000000	775.0000000	545.0000000
745.0000000	930.0000000	810.0000000	490.0000000
780.0000000	980.0000000	820.0000000	495.0000000
725.0000000	990.0000000	770.0000000	420.0000000
805.0000000	985.0000000	765.0000000	480.0000000
745.0000000	915.0000000	820.0000000	490.0000000
715.0000000	945.0000000	810.0000000	540.0000000
690.0000000	910.0000000	700.0000000	510.0000000
770.0000000	930.0000000	815.0000000	520.0000000
780.0000000	890.0000000	860.0000000	480.0000000
695.0000000	895.0000000	775.0000000	440.0000000
815.0000000	910.0000000	835.0000000	515.0000000
755.0000000	950.0000000	830.0000000	535.0000000
745.0000000	900.0000000	840.0000000	495.0000000
710.0000000	860.0000000	720.0000000	440.0000000
820.0000000	1020.0000000	845.0000000	525.0000000
745.0000000	920.0000000	810.0000000	470.0000000
800.0000000	1055.0000000	850.0000000	520.0000000
730.0000000	955.0000000	805.0000000	515.0000000
870.0000000	1105.0000000	940.0000000	500.0000000
770.0000000	995.0000000	825.0000000	540.0000000
835.0000000	1085.0000000	930.0000000	500.0000000
785.0000000	985.0000000	805.0000000	450.0000000
735.0000000	900.0000000	855.0000000	530.0000000
680.0000000	910.0000000	775.0000000	540.0000000

760.0000000	965.0000000	770.0000000	575.0000000
745.0000000	965.0000000	805.0000000	530.0000000
760.0000000	990.0000000	820.0000000	535.0000000
720.0000000	925.0000000	835.0000000	525.0000000
785.0000000	925.0000000	860.0000000	510.0000000
775.0000000	950.0000000	850.0000000	540.0000000
725.0000000	880.0000000	800.0000000	430.0000000
770.0000000	980.0000000	865.0000000	550.0000000
680.0000000	860.0000000	790.0000000	465.0000000
720.0000000	900.0000000	765.0000000	510.0000000
745.0000000	915.0000000	855.0000000	530.0000000
745.0000000	930.0000000	840.0000000	505.0000000
760.0000000	920.0000000	825.0000000	510.0000000
760.0000000	915.0000000	790.0000000	505.0000000
760.0000000	1015.0000000	805.0000000	470.0000000
750.0000000	995.0000000	840.0000000	515.0000000
755.0000000	920.0000000	885.0000000	590.0000000
750.0000000	910.0000000	860.0000000	485.0000000
760.0000000	925.0000000	765.0000000	540.0000000
770.0000000	940.0000000	865.0000000	510.0000000
765.0000000	900.0000000	815.0000000	560.0000000
715.0000000	870.0000000	820.0000000	530.0000000
710.0000000	920.0000000	770.0000000	490.0000000
790.0000000	950.0000000	865.0000000	515.0000000
770.0000000	870.0000000	950.0000000	680.0000000
780.0000000	1025.0000000	865.0000000	570.0000000
760.0000000	960.0000000	860.0000000	540.0000000
685.0000000	870.0000000	750.0000000	485.0000000
780.0000000	995.0000000	860.0000000	600.0000000
755.0000000	980.0000000	850.0000000	505.0000000
730.0000000	940.0000000	835.0000000	555.0000000
790.0000000	1030.0000000	920.0000000	525.0000000
740.0000000	955.0000000	870.0000000	505.0000000
760.0000000	990.0000000	860.0000000	550.0000000
745.0000000	955.0000000	815.0000000	545.0000000
750.0000000	955.0000000	805.0000000	540.0000000
745.0000000	985.0000000	810.0000000	610.0000000
735.0000000	945.0000000	810.0000000	480.0000000
815.0000000	1060.0000000	860.0000000	580.0000000
740.0000000	925.0000000	850.0000000	535.0000000
730.0000000	940.0000000	760.0000000	510.0000000
740.0000000	935.0000000	830.0000000	520.0000000
670.0000000	865.0000000	710.0000000	470.0000000
725.0000000	885.0000000	815.0000000	500.0000000
690.0000000	880.0000000	790.0000000	490.0000000
685.0000000	875.0000000	810.0000000	510.0000000
710.0000000	870.0000000	760.0000000	480.0000000
670.0000000	850.0000000	785.0000000	445.0000000
755.0000000	960.0000000	845.0000000	445.0000000
745.0000000	920.0000000	840.0000000	470.0000000
650.0000000	800.0000000	695.0000000	440.0000000
760.0000000	900.0000000	835.0000000	445.0000000
700.0000000	905.0000000	790.0000000	515.0000000
670.0000000	850.0000000	735.0000000	520.0000000
765.0000000	950.0000000	850.0000000	525.0000000
675.0000000	850.0000000	750.0000000	485.0000000
730.0000000	880.0000000	795.0000000	495.0000000
680.0000000	865.0000000	760.0000000	580.0000000
765.0000000	945.0000000	805.0000000	555.0000000

DISTRIBUTION DETAILS FOR REFERENCE SAMPLE

MEANS FOR SAMPLE 1

748.47656
722.92969
625.35938
417.85156

STANDARD DEVIATIONS FOR SAMPLE

31.06736
36.68083
32.03166
40.66709

MEANS FOR SAMPLE 2

746.15789
938.47368
818.15789
515.34211

STANDARD DEVIATIONS FOR SAMPLE

41.29198
57.92546
49.60510
43.00948

VARIABLE SKEWNESS T(SKEW) KURTOSIS T(KURT)

1	-0.2013	-0.9403	-0.1028	-0.2419
2	0.0007	0.0032	-0.1935	-0.4554
3	-0.2992	-1.3981	0.0725	0.1706
4	3.6562	17.0832	26.6019	62.6037

ST.DEV.FOR SKEWNESS = 0.2140261

ST.DEV.FOR KURTOSIS = 0.4249253

DISTRIBUTION DETAILS FOR COMPARISON SAMPLE

VARIABLE	SKEWNESS	T(SKEW)	KURTOSIS	T(KURT)
1	0.0878	0.3549	0.0946	0.1930
2	0.5615	2.2690	0.3348	0.6830
3	-0.0905	-0.3658	0.2707	0.5522
4	0.4673	1.8882	1.2869	2.6254

ST.DEV.FOR SKEWNESS = 0.2474635

ST.DEV.FOR KURTOSIS = 0.4901701

ROW COVARIANCE MATRIX 1 ( REFERENCE MATRIX )

	1	2	3	4
1	965.18055	778.90779	551.71321	327.60673
2	778.90779	1345.48321	709.66720	278.98161
3	551.71321	709.66720	1026.02731	282.37266
4	327.60673	278.98161	282.37266	1653.81244

ROW COVARIANCE MATRIX 2 ( COMPARISON MATRIX )

	1	2	3	4
1	1755.02800	1889.81803	1490.71949	454.33371
2	1889.81803	3355.35836	1623.48824	548.37346
3	1490.71949	1623.48824	2460.66630	706.88691
4	454.33371	548.37346	706.88691	1849.81523

ROW POOLED MATRIX

	1	2	3	4
1	1279,86679	1251,42165	951,10955	381,50871
2	1251,42165	2200,36223	1098,35126	393,56457
3	951,10955	1098,35126	1636,23575	462,93528
4	381,50871	393,56457	462,93528	1737,18014

\*\*\*\*\*

\*\*\*\*\*

EIGENVALUES FOR REFERENCE MATRIX  
 2740,99516 1423,34108 479,04893 347,11834

EIGENVECTORS FOR REFERENCE MATRIX

	1	2	3	4
1	-0,48781	0,17283	0,35113	0,78030
2	-0,59810	0,33807	0,37958	-0,61960
3	-0,47418	0,20227	-0,85575	0,04385
4	-0,42363	-0,90273	0,01763	-0,07282

\*\*\*\*\*

EIGENVALUES FOR COMPARISON MATRIX  
 6176,19985 1725,86087 1129,30396 339,50320

EIGENVECTORS FOR COMPARISON MATRIX

	1	2	3	4
1	-0,47962	-0,11076	0,08636	0,86616
2	-0,66687	-0,38264	-0,52438	-0,36591
3	-0,52583	0,21152	0,75090	-0,33898
4	-0,22081	0,89251	-0,39206	0,03095

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EIGENVALUES FOR POOLED MATRIX  
 4175,13456 1552,62951 765,53415 360,34669

EIGENVECTORS FOR POOLED MATRIX

	1	2	3	4
1	-0,48626	0,12064	0,01965	0,86522
2	-0,64950	0,29267	0,56313	-0,41865
3	-0,51393	0,02954	-0,81219	-0,27450
4	-0,27853	-0,94805	0,15115	-0,02778

LOG DETERMINANTS  
 REFERENCE COMPARISON  
 27,198 29,039 POLED  
 28,212

B SQUARE = 51,0763999  
 BETA SQUARE = 0.2004864  
 DEGREES OF FREEDOM = 10

TRACE B = 9370.8678775

PERCENTAGES FOR B(I)  
 65.9085149 16.4173002 12.0512206 3.6229643

#### ORIENTATION OF ELLIPSOIDS

VECTOR	CHISQUARE	DF
1	13.901	3
2	11.784	3
3	34.449	3
4	36.211	3

TEST = 23.2100000

#### STATISTICAL DISTANCE FOR HETEROGENEUS COVARIANCE MATRICES

NA = 128 NB = 95

#### ITERATIONS FOR DISTANCE CALCULATION SAMPLE DIFFERENCE MEAN VECTOR

-2.31867  
 215.54400  
 192.29852  
 97.99054

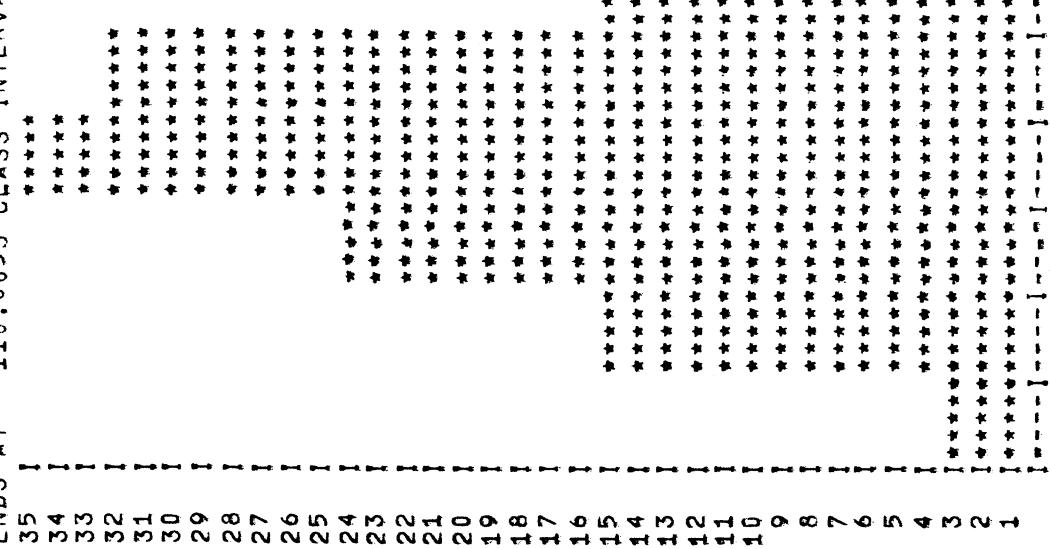
ITERATION	Z	D	DSC	DIFF T
0	-3.61526	8.792	77.304	0.0000000
0	36.21169	8.745	76.470	0.0000000
0	15.32475	8.784	77.164	0.0000000
0	5.67311	8.792	77.292	0.0000000
4	0.98963	8.793	77.311	0.0625000
5	-1.32195	8.793	77.311	0.0312500
6	-0.16852	8.793	77.312	0.0156250
7	0.40995	8.793	77.312	0.0078125
8	0.12056	8.793	77.312	0.0039063
9	-0.02402	8.793	77.312	0.0019531
10	0.04826	8.793	77.312	0.0009766

#### ESTIMATE OF DISCRIMINANT VECTOR

-0.32796  
 0.19091  
 0.16544  
 0.03858

HISTOGRAM FOR DISCRIMINANTS

INTERVAL = 5.9571 DATA STARTS AT •9.1373  
ENDS AT 110.0053 CLASS INTERVALS = 20



HETEROGENEOUS DISTANCE BY AVERAGING COVARIANCE MATRICES

ROW COVARIANCE MATRIX

	1	2	3	4
1	1335.10427	1334.36291	1021.21635	390.97022
2	1334.36291	2350.42078	1166.57772	413.67753
3	1021.21635	1166.57772	1743.34680	494.62978
4	390.97022	413.67753	494.62978	1751.81383

DISTANCE, DISTANCE SQUARED  
8.783            77.134

HETEROGENEOUS SIGNIFICANCES

MULTIPLICATION FACTOR

AN = 0.8615030

ITERATION REYMENTS DSQ TEST

1	83.7096
2	68.4095
3	74.0644
4	81.8624
5	81.0875
6	85.7856
7	82.6404
8	81.9864
9	71.4774
10	87.0855
11	68.8185
12	97.1551
13	90.7586
14	79.9636
15	78.6520
16	76.5856
17	100.2027
18	97.7471
19	84.5251
20	98.4155
21	88.5189
22	97.8546
23	90.6790
24	81.1956
25	91.8051
26	80.7465
27	77.5857
28	92.5904
29	93.7228
30	131.5444

SIGNIFICANCE COMPUTATIONS

AVE,DSQ            TSQ  
83.780            3979.543

VARIANCE RATIO IS = 963.13 FOR 94 AND 4 DEGREES OF FREEDOM

STATISTICAL DISTANCE

VAR	N1 = 128 MEAN VECTOR 1	N2 = 95 MEAN VECTOR 2	DF = 221 DIFF.VECTOR
1	748.477	-0.328	-2.319
2	722.930	0.191	215.544
3	625.859	0.165	192.299
4	417.852	0.039	97.991

DISCRIMINATOR COEFFICIENTS  
-0.32631 0.19354 0.16597 0.03999

RESULTS FOR MAHANALOBIS DSQUARE

D	D SQUARE
8.84919	78.30818

SIGNIFICANCE FOR D SQUARE AND T SQUARE

T SQUARE = 4270.078

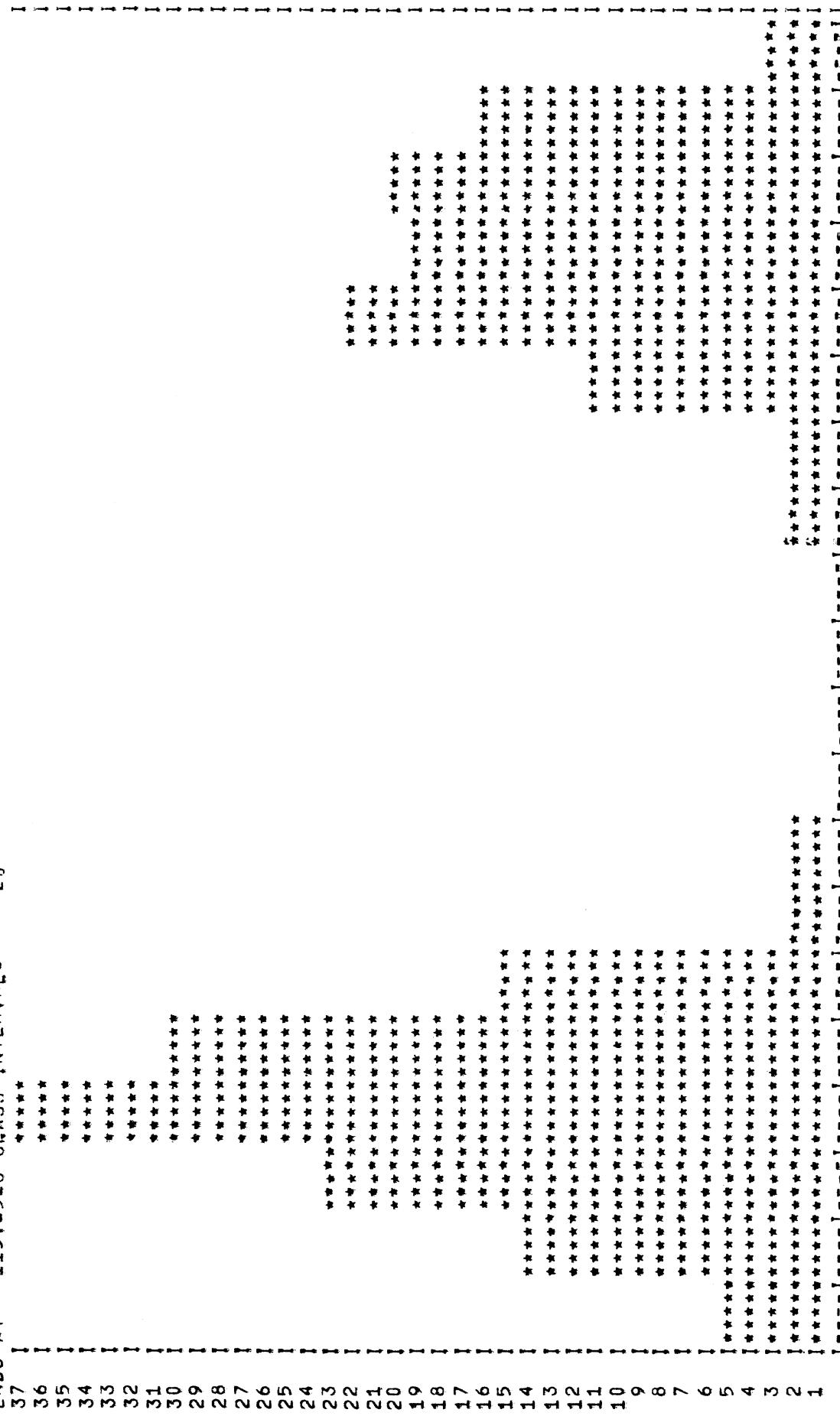
F = 1053.03

DF1 = 4 DF2 = 218

HISTOGRAM FOR DISCRIMINANTS

INTERVAL = 6.0133 DATA STARTS AT -5.0137

ENDS AT 115.2513 CLASS INTERVALS = 20



KANSAS GEOLOGICAL SURVEY COMPUTER PROGRAM  
THE UNIVERSITY OF KANSAS, LAWRENCE

PROGRAM ABSTRACT

Title (If subroutine state in title):

FORTRAN IV program for generalized statistical distance and analysis of covariance  
matrices for the CDC 3600 computer.

Date: Final version in March 1969

Author, organization: R.A. Reymert and Hans-Ake Ramden (University of Uppsala), and Warren J.  
Wahlstedt (Kansas Geological Survey)

Direct inquiries to: Authors or

Name: D.F. Merriam Address: Kansas Geological Survey  
University of Kansas, Lawrence, Kansas

Purpose/description: For computing generalized distances and discriminant functions in the situations  
of homogeneous or heterogeneous covariance matrices and tests for homogeneity of covariance  
matrices.

Mathematical method: Generalized distances by the methods of Mahalanobis, Anderson and Bahadur,  
and Reymert.

Restrictions, range: The programs accept matrices up to 45 x 45

Computer manufacturer: CDC Model: 3600

Programming language:

Memory required: 32 K Approximate running time:

Special peripheral equipment required:

Remarks (special compilers or operating systems, required word lengths, number of successful runs, other machine versions, additional information useful for operation or modification of program)

Subroutine REYDST requires access to a random access unit to be efficient. Running  
time of the example was about 15 minutes.













(continued from inside front cover)

9.	FORTRAN IV program for mathematical simulation of marine sedimentation with IBM 7040 or 7094 computers, by J.W. Harbaugh and W.J. Wahlstedt, 1967 . . . . .	\$1.00
10.	Three-dimensional response surface program in FORTRAN II for the IBM 1620 computer, by R.J. Sampson and J.C. Davis, 1967 . . . . .	\$0.75
11.	FORTRAN IV program for vector trend analyses of directional data, by W.T. Fox, 1967 . . . . .	\$1.00
12.	Computer applications in the earth sciences: Colloquium on trend analysis, edited by D.F. Merriam and N.C. Cocke, 1967 . . . . .	\$1.00
13.	FORTRAN IV computer programs for Markov chain experiments in geology, by W.C. Krumbein, 1967 . . . . .	\$1.00
14.	FORTRAN IV programs to determine surface roughness in topography for the CDC 3400 computer, by R.D. Hobson, 1967 . . . . .	\$1.00
15.	FORTRAN II program for progressive linear fit of surfaces on a quadratic base using an IBM 1620 computer, by A.J. Cole, C. Jordan, and D.F. Merriam, 1967 . . . . .	\$1.00
16.	FORTRAN IV program for the GE 625 to compute the power spectrum of geological surfaces, by J.E. Esler and F.W. Preston, 1967 . . . . .	\$0.75
17.	FORTRAN IV program for Q-mode cluster analysis of nonquantitative data using IBM 7090/7094 computers, by G.F. Bonham-Carter, 1967 . . . . .	\$1.00
18.	Computer applications in the earth sciences: Colloquium on time-series analysis, D.F. Merriam, editor, 1967 . . . . .	\$1.00
19.	FORTRAN II time-trend package for the IBM 1620 computer, by J.C. Davis and R.J. Sampson, 1967 . . . . .	\$1.00
20.	Computer programs for multivariate analysis in geology, D.F. Merriam, editor, 1968 . . . . .	\$1.00
21.	FORTRAN IV program for computation and display of principal components, by W.J. Wahlstedt and J.C. Davis, 1968 . . . . .	\$1.00
22.	Computer applications in the earth sciences: Colloquium on simulation, D.F. Merriam and N.C. Cocke, editors, 1968 . . . . .	\$1.00
23.	Computer programs for automatic contouring, by D.B. McIntyre, D.D. Pollard, and R. Smith, 1968 . . . . .	\$1.50
24.	Mathematical model and FORTRAN IV program for computer simulation of deltaic sedimentation, by G.F. Bonham-Carter and A.J. Sutherland, 1968 . . . . .	\$1.00
25.	FORTRAN IV CDC 6400 computer program for analysis of subsurface fold geometry, by E.H.T. Whitten, 1968 . . . . .	\$1.00
26.	FORTRAN IV computer program for simulation of transgression and regression with continuous-time Markov models, by W.C. Krumbein, 1968 . . . . .	\$1.00
27.	Stepwise regression and nonpolynomial models in trend analysis, by A.T. Miesch and J.J. Connor, 1968 . . . . .	\$1.00
28.	KWIKR8 a FORTRAN IV program for multiple regression and geologic trend analysis, by J.E. Esler, P.F. Smith, and J.C. Davis, 1968 . . . . .	\$1.00
29.	FORTRAN IV program for harmonic trend analysis using double Fourier series and regularly gridded data for the GE 625 computer, by J.W. Harbaugh and M.J. Sackin, 1968 . . . . .	\$1.00
30.	Sampling a geological population (workshop on experiment in sampling), by J.C. Griffiths and C.W. Ondrick, 1968 . . . . .	\$1.00
31.	Multivariate procedures and FORTRAN IV program for evaluation and improvement of classifications, by Ferruh Demirmen, 1969 . . . . .	\$1.00
32.	FORTRAN IV programs for canonical correlation and canonical trend-surface analysis, by P.J. Lee, 1969 . . . . .	\$1.00
33.	FORTRAN IV program for construction of Pi diagrams with the Univac 1108 computer, by Jeffrey Warner, 1969 . . . . .	\$1.00
34.	FORTRAN IV program for nonlinear estimation, by R.B. McCammon, 1969 . . . . .	\$0.75
35.	FORTRAN IV computer program for fitting observed count data to discrete distribution models of binomial, Poisson and negative binomial, by C.W. Ondrick and J.C. Griffiths, 1969 . . . . .	\$0.75
36.	GRAFPAC, graphic output subroutines for the GE 635 computer, by F.J. Rohlf, 1969 . . . . .	\$1.00
37.	An iterative approach to the fitting of trend surfaces, by A.J. Cole . . . . .	\$1.00
38.	FORTRAN II programs for 8 methods of cluster analysis (CLUSTAN I), by David Wishart, 1969 . . . . .	\$1.50
39.	FORTRAN IV program for the generalized statistical distance and analysis of covariance matrices for the CDC 3600 computer, by R.A. Reymont, Hans-Ake Ramden, and Warren J. Wahlstedt . . . . .	\$1.00

